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CEO political ideology and voluntary forward-looking disclosure

Abstract

This study investigates whether and, if so, how the information disclosure preferences differ systematically between Republican CEOs and Democrat CEOs in the context of management earnings forecasts. We find that Republican CEOs prefer a less asymmetric information environment than Democrat CEOs, and thus make more frequent, timelier, and more accurate disclosures than Democrat CEOs. Results using the propensity score matched sample and difference-in-differences analysis show that our results are unlikely to be driven by potential endogeneity. Our results are robust to controlling for various CEO characteristics and are stronger for firms with higher levels of institutional ownership and litigation risk.

Keywords: CEO Political Ideology; Management Earnings Forecasts; Conservatism; Information disclosure; Republican versus Democrat.

JEL classification: D82, M41, P16, G34, G40.

1. Introduction¹

At present, the U.S. is experiencing an unprecedented degree of political polarization. The Pew Research Center documents that over the period of 1994 to 2017, the average partisan gap (the difference in opinion between supporters of Republicans and Democrats) has increased from 15 to 36 percentage points.² According to a recent Gallup report, political identity influences people's views on a wide variety of matters that are often not directly related to politics (Newport, 2019). Political ideology seems to affect a broad spectrum of our life choices, ranging from what we eat to our perception of climate change. Consistent with these findings, a recent Forbes report argues that, for many, political ideology is becoming an official religion.³

A growing body of research has investigated the effect of managers' political ideology on their corporate policies. This literature shows that Republican CEOs have more conservative investment and financial policies than Democrats (e.g., Hong and Kostovetsky, 2012; Di Giuli and Kostovetsky, 2014; Hutton et al., 2014; Francis et al., 2016; Elnahas and Kim, 2017), are less likely to engage in earnings management, pay lower audit fees, and have higher financial reporting quality (Dong et al., 2018). These findings are in line with predictions of the behavioral consistency principle.^{4,5} However, one may argue that the actual effect of CEO political ideology might be undermined in these studies, because financial and investment

¹ Abbreviations used in this paper:

MEFs – management earnings forecasts

² A summary of Pew's report can be found at the following link: https://www.people-press.org/2017/10/05/the-partisan-divide-on-political-values-grows-even-wider/

³ See more details at https://www.forbes.com/sites/johnhart/2017/11/30/is-ideology-becoming-americas-official-religion/#ce0893a164b3

⁴ Similarly, Wintoki and Xi (2019) document that mutual fund managers allocate assets to firms whose executives and directors share a similar political partian affiliation.

⁵ Researchers study CEO political ideology, as a personal trait, because personal political ideology is established in early adulthood and becomes relatively consistent over time (Jost and Amodio, 2012). Further, political ideology can be clearly measured based on a CEO's political donations and hence is subject to little measurement error.

policies tend to be persistent, and thus, are typically subject to limited managerial discretion (Fee et al., 2013).⁶

Management earnings forecasts (MEFs), one of voluntary disclosures over which managers have a higher degree of managerial discretion (e.g., Houston et al., 2010; Cheng et al., 2013), could provide a much cleaner setting in which to investigate the impact of CEO political ideology on corporate policy choices.⁷ This is so, in particular, because managers can exercise their full discretion over management earnings forecasts (MEFs) to alter investor expectations about the future stock price, mitigate information asymmetry (Brown and Hillegeist, 2007), reduce the cost of capital (Baginski and Rakow, 2012), increase analyst following (Ajinkya et al., 2005), and enhance a firm's reputation for accurate and transparent reporting (Graham et al., 2005). In this study, we investigate a hitherto under-explored question of whether and, if so, how CEO political ideology influences MEFs.

Another reason that makes this study different from prior studies on CEO political ideology and makes its research question worthy of investigation is that transparency and quality disclosure are not exclusively claimed by one political party. Prior studies find that Republican CEOs tend to adopt more conservative financial and investment policies (Hong and Kostovetsky, 2012; Hutton et al., 2014; Francis et al., 2016; Elnahas and Kim, 2017), while Democrat CEOs tend to engage more in corporate social responsibility (CSR) activities (Di Giuli and Kostovetsky, 2014). This finding enhances our understanding of how CEOs personal traits affect their firms' decision

⁶ Several empirical studies show that firms adjust their capital structure slowly over multiple years (e.g., Flannery and Rangan, 2006).

⁷ Management earnings forecasts are defined as voluntary managerial disclosures predicting earnings prior to the actual earnings reporting date.

making. However, it is not surprising that conservative CEOs act conservatively and pro-social CEOs engage more in socially responsible activities.

However. who. in practice, advocates for transparency and quality disclosure—conservatives or liberals? Berliner (2014) sheds some light on this matter by investigating the political origins of transparency through tracing the emergence of the Freedom of Information Acts (FOIA) globally. He shows that while the FOIA in Canada was passed in 1982 by the Liberal party, it was opposed and delayed for years by the Liberal Democratic Party (LDP) in Japan. In the U.S., the FOIA was first proposed by the Democrat congressman John Moss. However, in the meantime, it was almost vetoed by the Democrat president Lyndon Johnson and was opposed by almost all federal agencies and departments. Recently, Joshua et al. (2017), in their Washington Post article, reviewed the results of a survey, conducted by a group of researchers at the University of Massachusetts, on how Republicans and Democrats support the basic democratic freedoms. Interestingly, Republicans and Democrats both say they support freedoms related to transparency and disclosure, like freedom of speech and freedom of the press. Given this background, we aim to provide systematic evidence and useful insights that help to better understand how CEOs' personal traits affect their firms' decision-making and contribute to the ongoing debate on who, Republicans or Democrats, favors more transparency and higher-quality disclosure, when in power.

Conservatism is defined by Wilson (1973: p.4) as "resistance to change and the tendency to prefer safe, traditional, and conventional forms of institutions and behavior." Since the early 1950s, political conservatism has been studied by political scientists, historians, sociologists, and philosophers, among others. During these decades of research, several theoretical frameworks have emerged to explain the psychology of politically conservative individuals (Jost et al., 2003). First, personality theories associate political conservatism with authoritarianism and intolerance of ambiguity (Peterson et al., 1993). Second, epistemic and existential need theories postulate that conservatives have a higher need for closure, desire for security and stability, and preference for the avoidance of threats and change (Jost et al., 1999). Finally, sociopolitical theories argue that conservatives have a higher preference for social dominance and system justification (Sidanius and Pratto, 1999).

These theories have interesting implications with regard to conservative CEOs' attitudes towards transparency and voluntary disclosure like MEFs. On the one hand, by definition, individuals with a high need for closure do not have a high preference for information disclosure. Further, failure of actual earnings to meet MEF could increase litigation risk (Francis et al., 1994) as well as CEO turnover (Lee et al., 2012). Consequently, transparency and high-quality disclosure can represent a threat to individuals with authoritarian personalities. As a result, the authoritarian nature and need for closure of politically conservative CEOs can foster their tendency to seize on information (Jost et al., 2003), and thus lead them to prefer less transparent disclosure. In this study, this effect of conservative political ideology on CEO disclosure policies is conveniently called the *authoritarian effect*.

On the other hand, prior research shows that high-quality MEFs have several significant benefits for firms as well as CEOs. At the firm level, high-quality MEFs increase firm value and reduce firm risk (Trueman, 1986), information asymmetry (Brown and Hillegeist, 2007), share price volatility (Graham et al., 2005), and the likelihood of litigation (Skinner, 1994). At the CEO level, high-quality MEFs enhance managerial reputation (Graham et al., 2005) and reduce career penalties in the form of bonus cuts, fewer stock grants, and forced turnover (Lee et al., 2012). These potential personal costs, which can be mitigated by high-quality MEFs, represent an important form of potential losses to CEOs. As a result, politically conservative CEOs' intolerance of ambiguity, desire for security (including job and financial security), and preference for the avoidance of uncertainty and threats can lead them to adopt more transparent and higher-quality disclosure policies. This effect is conveniently called the *precautionary effect*.

Hence, Republican CEOs' attitudes towards MEFs are determined by the trade-off between the two effects mentioned above, that is: (i) the perceived benefits achieved by satisfying their authoritarian needs through seizing on information (the authoritarian effect); and (ii) the perceived losses prevented by adopting more transparent disclosure policies (the precautionary effect). Political conservatives have been described by personality theoreticians as generally more sensitive to the threat of loss and to negatively framed outcomes (e.g., potential losses) than to positively framed outcomes (e.g., potential gains) (Jost et al., 2003). Accordingly, we conjecture that Republican CEOs are more motivated by the *precautionary effect* than by the *authoritarian effect*, leading them to adopt more transparent disclosure policies.

Conservative CEOs' preference for more transparent MEFs was apparent when Hewlett Packard's (HP) Democrat CEO Lewis E. Platt was succeeded by the renowned Republican Carly Fiorina in 1999. HP's MEFs experienced a drastic change upon this move from a Democrat CEO to a Republican CEO. Specifically, whereas Mr. Platt had an average forecast issuance, frequency, and accuracy of 0.143, 1.00, and 1.00, respectively, Mrs. Fiorina had significantly higher forecast issuance, frequency, and accuracy of 0.60, 3.33, and 2.43, respectively.⁸ In this paper, we present evidence that HP is not a unique example; instead, it is just the tip of the iceberg.

⁸ Forecast issuance is a binary variable that captures the likelihood of issuing MEFs, forecast frequency refers to the number of MEF issues during the fiscal year, and forecast accuracy captures the difference between MEF and actual earnings. Appendix A provides operational definitions of the three variables.

Following Hutton et al. (2014), among others, we use CEOs' personal political donations to a candidate or a party committee to measure their political ideology.⁹ Using a sample covering the period of 1993–2016, we examine the effect of CEO political ideology on managers' forecast preferences and various properties of MEFs, including: (i) managers' preference for forecast issuance and frequency; (ii) managers' preference for forecast horizon and range; (iii) MEF credibility, including forecast bias and accuracy, the likelihood of the company's earnings missing its own forecasts, and the likelihood of releasing bad versus good news; and (iv) MEF surprises, referring to the likelihood of having a positive, negative, or neutral earnings surprise.

To provide systematic evidence on the impact of CEO political ideology on CEO voluntary disclosure strategy, we first examine whether, and if so, how, CEO political ideology influences the likelihood and frequency of issuing MEFs. In so doing, we classify CEOs into Republican CEOs and Democrat CEOs, using data on CEOs' donations to two political parties. We find that, on average, Republican CEOs are approximately 13 percent more likely to issue forecasts than Democrat CEOs. Further, among CEOs who issue MEFs, Republican CEOs disclose 16.5 percent more forecasts, on average, compared to Democrat CEOs. Second, we test the effect of CEOs' political ideology on the earnings forecasts horizon and forecast range. Due to their preference for avoiding a negative market reaction and litigation if they miss their own forecasts, we expect Republican CEOs to prefer range over point forecasts, compared to Democrat CEOs. Further, due to their preference for avoiding ambiguity and information asymmetry, we expect that Republican managers issue timelier forecasts than Democrat CEOs. Our results show that Republican CEOs are more likely to issue range forecasts and issue forecasts in a timelier fashion than Democrat

⁹ Other studies adopting similar measures include Hong and Kostovetsky (2012), Francis et al. (2016), Elnahas and Kim (2017), and Bhandari et al. (2018).

CEOs, which is consistent with their conservative political ideology (Hutton et al., 2014). For instance, Republican CEOs issue 12.7 percent more range estimates than non-Republican CEOs. Further, Republican CEOs have an average forecast horizon that is approximately 11 percent longer than that of Democrat CEOs.¹⁰

Third, we test for the effect of CEO political ideology on the credibility of MEFs. We find that Republican CEOs are less likely to miss their own forecasts and have a lower degree of forecasting bias than Democrat CEOs. For instance, on average, Republican CEOs are 14.3 percent less likely to issue optimistically biased forecasts and 14.2 percent less likely to miss forecasts in any given year, compared to Democrat CEOs. We find further evidence of a positive relation between Republican CEO ideology and forecast accuracy. For instance, forecasts made by Republican CEOs are, on average, 8.7 percent more accurate than those made by CEOs with other political ideologies.

Skinner (1994) argues that managers disclose bad news forecasts to deter future litigation. Hutton et al. (2015) find that Republican CEOs are less likely than other CEOs to be the subject of litigation related to securities fraud. Our results show that Republican CEOs are more likely to issue bad news; they disclose bad news more frequently and less likely to hoard bad news compared to good news forecast. Finally, we test the effect of CEO political ideology on earnings surprises. Skinner and Sloan (2002) argue that small negative earnings surprises could cause a significant stock price decline, whereas Kasznik and McNichols (2002) document that stock prices increase significantly due to small positive earnings surprises. Our evidence shows that firms with Republican CEOs are more likely to have positive earnings surprises than negative earnings

¹⁰ We use the terms, Republican (Democratic) CEOs and conservative (liberal) CEOs interchangeably throughout the paper.

surprises. Our results are robust to controlling for other determinants of MEFs. Overall, our study provides evidence that Republican CEOs are more likely than other CEOs to disseminate private information and prefer an environment of less ambiguity and lower information asymmetry.

We conduct several tests to alleviate concerns about potential endogeneity in our baseline results. First, we use the propensity score matching (PSM) technique and difference-in-differences (DID) regressions around CEO turnover events to address the possibility that certain types of firms and industries hire CEOs with a similar political ideology to implement their desired corporate policies. Second, we address the possibility that our baseline results are affected by correlated omitted CEO characteristics, incentives, or overconfidence. Specifically, we control for CEO payperformance sensitivity (Delta) and CEO risk-taking incentives (Vega), CEO tenure, CEO gender, role duality, CEO age, and CEO overconfidence. Third, to address possible measurement error inherent in our baseline measures of CEO political ideology (for which we follow Bhandari et al. (2018) and Hong and Kostovetsky (2012)), we construct alternative measures of CEO political ideology, similar to Hutton et al. (2014) and Elnahas and Kim (2017). We find that our results are robust to these alternative measures. Overall, we find that our baseline results are robust to the use of PSM, DID design, and other sensitivity checks, suggesting that our results are unlikely to be driven by potential endogeneity caused by correlated omitted variables, reverse causality, and/or errors in measurement.

To further test our conjecture that Republican CEOs favor higher-quality MEFs to avoid litigation and other disciplinary actions (i.e., the precautionary effect), we also conduct crosssectional tests using subsamples of firms with large and small institutional ownership and those with high and low litigation risk. The results of these subsample tests lend strong support to the presence of the precautionary effect or the precautionary effect dominating the authoritarian effect. Specifically, in firms with a high level of institutional ownership, Republican CEOs are 17.7 percent more likely to issue forecasts than non-Republican CEOs; in contrast, they are only 9.7 percent more likely to do so in firms with a low level of institutional ownership. Similarly, in firms with high litigation risk, the accuracy of forecasts issued by Republican CEOs is 20.6 percent higher than those issued by non-Republican CEOs, whereas it is only 5.4 percent higher in firms with low litigation risk.

We perform a battery of additional robustness tests. To address the coverage issue of the management guidance database, we exclude firms that have never issued earnings forecasts during our sample period and find similar results to our baseline regression.¹¹ Further, we test the robustness of our results to the use of several alternative measures of CEO political ideology and the inclusion of additional control variables such as CEO overconfidence. We find that our main results are robust to these sensitivity checks.

This study makes several contributions to the literature. First, in a broad sense, our study contributes to the literature on the effects of cultural traits on corporate decision-making (e.g., Ahern et al., 2015). More specifically, it contributes to the recent stream of research investigating the effect of CEO political ideology on corporate policy choices. This strand of research focuses largely on corporate investment and financial policies. For instance, Hutton et al. (2014) find that Republican managers pursue more conservative corporate policies, such as lower debt, lower R&D expenditure, and less risky investment policies. Similarly, Elnahas and Kim (2017) find that Republican CEOs are less likely to engage in M&A activities and avoid diversification. Francis et al. (2016) find that Republican managers are less likely to engage in corporate tax avoidance. Hong

¹¹ Furthermore, we run a robustness test that includes only firm-year observations in which a firm issues one or more MEFs, and again find support for our main findings. These results are available in the Internet Appendix.

and Kostovetsky (2012) find that mutual fund managers who make political donations to Democrats are less likely to invest in socially irresponsible firms. Surprisingly, however, this line of research has paid little attention to the role of CEO political ideology in shaping corporate disclosure policies. To the best of our knowledge, our study is the first to examine the impact of CEO political ideology on voluntary disclosure in the form of management earnings guidance or MEFs. Our study fills this knowledge gap by providing large-sample, systematic evidence on the relation between CEO political ideology and a firm's decisions on whether, when, and how to issue forward-looking earnings guidance.

Second, we note that studying the impact of CEO political ideology on MEFs is less subject to confounding effects than studying its impact on investment and financial policies. This is because investment and financing policies have lower managerial discretion and thus tend to be persistent (Fee et al., 2013). In contrast, MEFs are voluntary and so are subject to a higher degree of managerial discretion (e.g., Houston et al., 2010; Cheng et al., 2013). Thus, MEFs provide a cleaner setting in which to investigate how CEOs' personality traits such as political ideology translate into their corporate policy choices, allowing us to make stronger and more reliable inferences on the role of CEO political ideology in shaping corporate policies.

Finally, this study contributes to the literature on understanding the determinants of MEFs and the association between CEO personal characteristics and MEFs. Hribar and Yang (2016) find that CEO overconfidence increases forecast issuance and optimism and reduces forecast precision. Similarly, Bamber et al. (2010) find that managers with finance and accounting backgrounds and those with military experience issue conservative earnings forecasts and prefer a more precise disclosure style. Further, Jiang et al. (2016) find that conservative analysts issue more frequent and accurate forecasts and produce better quality research. Our study extends this literature and

presents evidence that Republican CEOs tend to issue more frequent, more accurate, and less biased forecasts.

The remainder of this paper proceeds as follows. Section 2 reviews the literature on behavioral consistency, CEO political ideology, and MEFs. Section 3 describes our data and the construction of our measures of CEO political ideology. Section 4 presents the empirical results and discusses their interpretation. Section 5 reports the robustness tests. The final section provides a summary and concluding remarks.

2. Literature review

Researchers in sociology and behavioral psychology have studied the implications of the behavioral consistency theory, and show that people behave consistently across different domains. For example, Epstein (1979) argues that individuals show stable behavioral patterns over time and across different areas. More recently, researchers in financial economics investigate whether the behavioral consistency theory can help understand various corporate decisions. For instance, Cronqvist et al. (2012) find consistent patterns between a firm's leverage decisions and the CEO's personal leverage decisions. Similarly, Biggerstaff et al. (2015) find that CEOs who are personally benefiting from options backdating tend to engage more in corporate misconduct.

Prior literature also suggests that CEOs' personal conservatism and risk-taking behavior are reflected in corporate decision-making. For instance, Graham et al. (2013) show that CEOs' optimism and risk aversion affect their firms' financial policies. Further, Benmelech and Frydman (2015) show that military CEOs are more conservative and behave more ethically than other CEOs. Similarly, Cain and McKeon (2016) argue that pilot CEOs are associated with higher equity return volatility, higher leverage, and higher levels of acquisition activity. Davidson et al. (2015) also find that CEOs' off-the-job behavior is positively related to their corporate behavior. In sum, managers' personality traits have been found to remain consistent across different domains and consequently influence their corporate policy choices.

2.1. Political ideology, personality traits, and corporate policies

Of late, researchers in corporate finance have paid considerable attention to CEO political ideology and its impact on corporate policy choices. Unlike other personal traits, a person's political orientation tends to be stable and consistent over time.¹² For example, Alford et al. (2005) argue that genes play a crucial role in shaping political attitudes, ideologies, and the strength of an individual's party identification. More importantly, Jost and Amodio (2012) argue that political ideology is established in early adulthood and becomes relatively consistent over time. The literature dealing with political ideology issues often substitutes "liberalism" and "conservatism" for "liberal" and "conservative," "left" and "right," or "Democrat" and "Republican."

Another reason why financial economists pay particular attention to CEO political ideology is the increasing polarization of the political environment in the U.S. This political environment has triggered several studies that aim to understand the differences between the two dominant political orientations in the U.S., namely Republicans and Democrats. Behavioral psychology literature finds stark ideological and psychological differences between the two groups; the main difference between the two political ideologies is the degree of openness to change. Jost et al. (2003) argue that conservatives avoid ambiguity, uncertainty, and complexity.

¹² If political ideology was subject to changes over time, then studying the relation between political ideologies into corporate policy choices would be potentially troublesome.

Relative to liberals, conservatives prefer to punish violators of social norms and prevent free riders, value job security more highly than task variety, fear losses, value financial security, value property rights, and show more respect for authority and preference for preserving the status quo. Further, Wilson (1973) shows that conservatives seek familiarity and safety and are resistant to change.

Several studies in financial economics investigate whether the above personality differences between conservatives and liberals translate into their firms' corporate decisions. These studies show that Republican managers prefer less risky investment and financial policies, are less likely to engage in mergers and acquisitions, avoid high information-asymmetry acquisitions, and are less likely to engage in corporate tax avoidance (Hutton et al., 2014; Francis et al., 2016; Elnahas and Kim, 2017). In contrast, Democrat managers are more likely to invest in corporate social responsibility and less likely to invest in socially irresponsible firms (Hong and Kostovetsky, 2012; Di Giuli and Kostovetsky, 2014). Furthermore, Hutton et al. (2015) find that Democrat managers are more likely to face litigation for securities fraud and intellectual property rights violations, whereas Republican managers are more likely to face civil rights, labor, and environmental litigation.¹³ More recently, Lee et al. (2018) show that Republican CEOs hold more outside directorship roles, regardless of the political regime. Finally, Babenko et al. (2020) find that CEO political orientation affects employee campaign contributions, whereby candidates supported by the CEO receive three times more contributions from employees than candidates not supported by the CEO.

¹³ Political ideology may affect the decisions of other decision makers too. For example, Jiang et al. (2016) argue that conservative analysts produce more accurate earnings forecasts, issue more frequent forecast updates, are less likely to deviate from benchmarks, and produce better quality research than other analysts.

In a broad sense, the aforementioned literature is mainly concerned about whether, and, if so, how CEO political ideology is associated with corporate investment and financial policies. Although CEOs certainly have a significant impact on firms' investment and financial policies, they do not have full autonomy over such policies due to several organizational considerations. For instance, Fee et al. (2013) argue that firms' investment and financing decisions are persistent and more likely to be determined by a firm's past policies and culture, and thus are subject to limited managerial discretion. Moreover, investment and financial policies that deviate greatly from value maximization are usually challenged by the market for corporate control.

In contrast, MEFs are voluntary, and managers have substantial discretion over whether, when, and how to issue earnings forecasts (Cheng et al., 2013). For instance, managers can temporarily stop issuing earnings forecasts if they fail to meet analysts' forecasts and resume issuance when they feel confident about meeting analyst forecasts (Houston et al., 2010). Similarly, managers tend to increase disclosure and bad news forecasts before insider purchases and equity offerings, strategically choose forecast precision, voluntarily disclose bad news forecasts, and tactically avoid negative earnings surprises (Skinner, 1994; Matsumoto, 2002; Cheng et al., 2013). Thus, MEFs provide an ideal setting in which to test how CEOs' personality traits affect corporate policy choices. CEOs, whether Republican or Democrats, may face several limitations in infusing their political ideologies into their firms' investment and financing policies. In contrast, CEOs, whether Republicans or Democrats, are better able to infuse their political ideologies into their voluntary disclosures, such as MEFs.

2.2. Management earnings forecasts

The extant literature identifies several firm-level and CEO-level characteristics as the main determinants of the likelihood of issuance, forecast frequency, and other properties of

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MEFs. At the firm level, MEFs depend on firms' legal and regulatory environment, investor demand, firm-specific litigation risk, earnings volatility, and managerial compensation incentives. For instance, investors tend to prefer investing in firms that have better disclosure policies and lower information asymmetry because such firms enjoy higher liquidity, lower cost of capital, and lower agency problems (e.g., Diamond and Verrecchia, 1991; Ajinkya et al., 2005). Moreover, firms with higher R&D expenditure are less likely to issue forecasts (Rogers and Stocken, 2005). Similarly, firms with higher earnings volatility tend to issue forecasts less often, and more profitable firms are likely to issue forecasts more frequently (Miller, 2002). Waymire (1985) argues that firms with more volatile earnings variability. Skinner (1994) argues that firms voluntarily disclose bad news forecasts to avoid subsequent litigation. Similarly, firms with higher ex-ante litigation risk and bad news are more likely to issue forecasts (Houston et al., 2019).

At the CEO level, researchers show that MEFs are affected by CEOs' compensation design, ability, overconfidence, and career concern. Stock-based incentives should increase MEF frequency and reduce agency problems in disclosure. For example, Baginski et al. (2018) find that managers' severance pay and stock option portfolios increase their earnings forecast accuracy. Baik et al. (2011) find that CEO ability is positively associated with forecast issuance, frequency, and accuracy. Further, Hribar and Yang (2016) argue that overconfident CEOs are more likely to issue MEFs. Their forecasts are more optimistic, and they are more likely to miss their own forecasts subsequently. Prior literature also recognizes the role that a CEO's career concern can play in shaping MEFs. Pae et al. (2016) find that CEOs with greater career concerns are more likely to provide downward earnings guidance and less likely to beat market

expectations; managers' career penalties, such as bonus cuts, fewer stock grants, and forced turnover, can also affect their earnings forecasts decisions. Moreover, Lee et al. (2012) find a positive relation between CEO turnovers and MEF errors.

Due to the high level of autonomy and discretion that CEOs have over voluntary disclosure, we expect disclosure to be affected by CEOs' personal preferences. As discussed earlier, we expect that conservative individuals such as Republican CEOs have less tolerance for ambiguity, uncertainty, and complexity, value job security, and have a higher fear of losses than other individuals. We further expect Republican CEOs to utilize MEFs as a mechanism to alter investors' earnings expectations, reduce future litigation concerns, and establish their reputation with regard to transparent and accurate reporting. Drawing on the above discussions, this study aims to provide large-sample evidence on the effect of CEO political ideology on various aspects of MEFs, including the likelihood of issuing forecast, forecast frequency, forecast ranges and horizon, and other properties of MEFs such as forecast bias, accuracy, forecast types, and credibility.

3. Data and sample selection

3.1. Data

We start with an initial sample of CEOs from the ExecuComp database covering firms in the S&P 1500 index from 1993 to 2016. We exclude financial firms (SIC between 6000 and 6999) and firms in the utilities industry (SIC between 4900 and 4999). Then, we merge CEOs' data with individual donations data obtained from the Federal Election Commission (FEC). The FEC publishes several types of files that identify donors who have made political contributions in amounts exceeding \$200. The individual's contribution files contain information on the contributor's name, city, state, zip code, employer, and occupation, as well as transaction date,

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amount, and unique committee ID. A committee is formed by a candidate or a political party to collect funds and contributions from individuals. The committee files contain a committee ID, name, type, party affiliation, city, state, zip code, and candidate ID.

CEOs and other corporate managers can contribute to political parties through their company's Political Action Committees (PACs) or directly by making individual contributions. Because PACs can contribute to multiple parties simultaneously (Cooper et al., 2010), we focus on individual political contributions to a candidate or a party committee for the purpose of measuring a CEO's political ideology. We identify the political contributions of CEOs using their contributions to Republican- and Democrat-affiliated Senate, House, presidential candidates, and political party committees.¹⁴ To identify a CEO's contributions to a political party, we link the contributor's name, occupation, employer, and transaction date provided by the FEC with the executive's name, company name, and fiscal year from the ExecuComp database.

Our management earnings per share (EPS) forecast data comes from I/B/E/S. We obtain actual earnings data from the I/B/E/S actuals file to ensure consistency between MEFs and EPS realization. Following Baik et al. (2011), Lee et al. (2012), and Hribar and Yang (2016), we exclude qualitative forecasts because we do not have well-defined criteria to identify whether such forecasts were missed. We also exclude earnings preannouncements (i.e., management forecasts that are issued after the fiscal year-end but before the actual earnings announcements (Ajinkya et

¹⁴ Details of the campaign contribution data are available at the Federal Election Commission (FEC), https://www.fec.gov//. We focus on the CEOs individual level campaign contribution rather than at the firm level for two reasons: (1) firms may contribute to exploit the political favors to maximize shareholders' benefits (Blau et al., 2013) whereas individuals' contributions mainly reflect their personal political preference; (2) to exploit political benefits, firms typically contribute to both parties and/or their contribution may vary depending on the congress majority in each election cycle, whereas individuals' contributions generally remain consistent across election cycles and they are mostly directed towards only one party.

al., 2005; Rogers and Stocken, 2005; Houston et al., 2019). Following prior literature, we restrict our analyses to annual EPS forecasts (Baik et al., 2011; Hribar and Yang, 2016).

Finally, we acquire data on firm-level characteristics from Compustat, stock return data from the Center for Research in Security Prices (CRSP), and institutional ownership data from Thomson-Reuters Institutional Holdings (13F). Combining these datasets results in a final sample of 33,951 unique firm-year observations for the period 1993–2016.

3.2. Measures of CEO political ideology

The association between CEO political ideology and corporate decisions has received considerable attention from recent studies, including Hong and Kostovetsky (2012), Hutton et al. (2014), Francis et al. (2016), Elnahas and Kim (2017), and Bhandari et al. (2018). These studies provide a variety of measures for a CEO's political ideology. We follow Bhandari et al. (2018) in constructing our first measure of the CEO' political ideology, *Rep_Dum*, which is an indicator variable that equals one if a CEO has donated more to the Republican Party than to the Democratic Party during her/his entire tenure, and zero otherwise. This is a long-term and robust measure of a CEO's political ideology, *Rep_Index* is similar to that in Hong and Kostovetsky (2012). Specifically, *Rep_Index* measures the percentage of a CEO's support for the Republican Party, calculated as the number of cycles in which a CEO donates exclusively to the Republican Party divided by her/his number of donation cycles in the sample period. This measure is based on the two-year election cycle, and a higher percentage shows stronger Republican affiliation.

To mitigate potential noises and biases inherent in specific measures of CEO political ideology and to ensure the comparability of our measures with those employed in the prior

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literature, we conduct a variety of robustness checks using several additional measures of CEO political ideology. Following Hutton et al. (2014), we use: (i) Rep dum_{cvcle}, which is an indicator variable that equals one if all of the donations made by a CEO in an election cycle are directed to the Republican Party (i.e., none to the Democratic Party), and zero otherwise; and (ii) *Rep* index_{cvcle}, which is calculated as a CEO's total donations to the Republican Party minus total donations to the Democratic Party divided by total donations to both parties in a given election cycle. Further, following Elnahas and Kim (2017), we use *Rep dum_{tenure}*, which is an indicator variable that equals one if all donations made by a CEO during her/his entire tenure are directed to the Republican Party, and zero otherwise. Finally, we check the robustness of our results to the use of two different measures of Democratic Party affiliation: (i) Dem Dum, which is an indicator variable that equals one if a CEO donated more to the Democratic Party than to the Republican Party during her/his entire tenure, and zero otherwise; and (ii) Dem Index, measures the percentage of a CEO's support for the Democratic Party, calculated as the number of cycles in which a CEO donates exclusively to the Democratic Party divided by her/his number of donation cycles in the sample period.¹⁵

3.3. Measures of voluntary disclosure

Following prior literature, we use several proxies to capture the likelihood of issuing MEFs as well as several of their different properties. First, to measure the likelihood of issuing MEFs, we use *Issue*, which is an indicator variable that equals one if a firm makes at least one annual earnings forecast in a fiscal year, and zero otherwise, and *Frequency*, which is the total number of annual earnings forecasts made by a firm in a fiscal year (Ajinkya et al., 2005; Baik et al., 2011;

¹⁵ Detailed descriptions of these variables are provided in Appendix A. Further, in addition to the measures of CEO political ideology reported in the paper, our Internet Appendix reports results using additional measures of Republican Party affiliation, Democratic Party affiliation, and political neutrality.

Houston et al., 2019). Second, we measure the forecast horizon, *Ln(Horizon)*, using the natural logarithm of one plus the average horizon of the annual earnings forecasts made by a firm in a fiscal year (Baik et al., 2011; Houston et al., 2019). For each forecast, the horizon is defined as the number of calendar days between the forecast announcement date and the corresponding period end date. We assign a value of zero when a firm makes no forecasts in a fiscal year. To measure the likelihood that a firm issues range instead of point forecasts, we use Range, which is an indicator variable that equals one if a firm issues range forecasts, and zero otherwise (Hribar and Yang, 2016). Forecast Miss is an indicator variable that equals one if a firm's earnings fall short of the forecast amount at least once in a year, and zero otherwise (Hribar and Yang, 2016). OptBias is an indicator variable that equals one if the average Bias in a year is positive, and zero otherwise (Ajinkya et al., 2005). Accuracy is the average forecast accuracy for all annual earnings forecasts made by a firm in a fiscal year (Houston et al., 2019). Next, we measure the nature of the news contained in MEFs. Bad News is an indicator variable that equals one if forecast news is negative, and zero otherwise.¹⁶ Good News is an indicator variable that equals one if forecast news is nonnegative, and zero otherwise (Rogers and Stocken, 2005; Rogers and Van Buskirk, 2013). Finally, we measure the direction of the earnings surprise. Positive Surprise is an indicator variable that equals one if an earnings surprise is greater than 0.0001, and zero otherwise. Negative Surprise is an indicator variable that equals one if an earnings surprise is less than -0.0001, and zero otherwise. *Neutral* Surprise is an indicator variable that equals one if an earnings surprise is between 0.0001 and -0.0001, and zero otherwise (Rogers and Van Buskirk, 2013).¹⁷

[Insert Table 1]

¹⁶ Forecast news is considered as negative (positive) if the management earnings forecast is less (greater) than the most recent mean analyst estimate.

¹⁷ More detailed descriptions of the calculation of these variables are provided in Appendix A.

Table 1 provides descriptive statistics for our measures of CEO political ideology, MEFs, and the control variables used in our baseline models. As shown in Table 1, the mean value of *Rep Dum* is 0.229, indicating that around 23 percent of CEOs make more contributions to the Republican Party than to the Democratic Party during their entire tenure. The mean value of Rep Index is 0.169, suggesting that in around 17 percent of cycles, CEOs exclusively donate to the Republican Party. These statistics are consistent with those reported by Bhandari et al. (2018), Hutton et al. (2014), and Hong and Kostovetsky (2012). The mean value of *Issue* is 0.35, which indicates that, on average, firms have a 35 percent likelihood of issuing at least one annual earnings forecast in a fiscal year. The mean value of *Frequency* is 1.55, which suggests that, on average, firms issue approximately 1.55 forecasts each fiscal year. The mean values of *Issue* and *Frequency* are comparable with those reported by Baik et al. (2011), Hribar and Yang (2016), and Houston et al. (2019). The mean value of Ln(Horizon) is 1.83, which means that, on average, firms in our sample release their earnings forecasts 68 days prior to the forecast period end date. The mean value of forecast accuracy is 1.04, which is comparable with that reported by Houston et al. (2019). Firms in our sample, on average, have a 19 percent probability of missing at least one earnings forecast, and 14.3 percent of sample firms in our sample issue optimistically biased forecasts. On average, 53 percent, 36.8 percent, and 10.2 percent of sample firms issue Positive Surprise, Negative Surprise, and Neutral Surprise forecasts, respectively, which is consistent with Rogers and Van Buskirk (2013). Further, Table 1 shows that, on average, institutional investors own about 54.2 percent of outstanding shares, 24 percent of firms are subject to increased risk of litigation, and 20.3 percent of firms have issued equity in the year. The mean values of firm characteristics are comparable with those reported in prior studies, including Ajinkya et al. (2005), Rogers and Stocken (2005), Baik et al. (2011), Hribar and Yang (2016), and Houston et al. (2019).

[Insert Table 2]

Table 2 presents Pearson correlation coefficients. As shown, we find a positive correlation between measures of Republican ideology and measures of the likelihood of MEF issuance (Issue and Frequency), indicating that Republican CEOs are more likely to share forward-looking information with the market, compared to non-Republican CEOs. Similarly, we find positive correlations between measures of Republican ideology and Accuracy, indicating that Republican CEOs make more accurate forecasts. We further find negative correlations between Republican ideology measures on the one hand and Forecast Miss and OptBias, on the other hand, indicating that Republican CEOs are less likely to miss earnings forecasts and that they make less biased forecasts, respectively. Consistent with prior research, Republican ideology is negatively correlated with RD and Volatility, and positively correlated with ROA (e.g., Hutton et al., 2014). In addition, we find a positive correlation of firm size with Issue, Frequency, Ln(Horizon), and Accuracy on the one hand, and a negative correlation of firm size with Forecast Miss and OptBias on the other hand. These correlations are consistent with the idea that larger firms issue more forecasts, have greater forecast accuracy, and have less forecasting bias (Ajinkya et al., 2005; Baik et al., 2011; Hribar and Yang, 2016; Houston et al., 2019).

4. Analysis and results

4.1. Baseline regression model

To formally test the association between CEO political ideology and MEFs, we estimate the following regression model:

$$MEF_{it} = \beta_0 + \beta_1 Republican_{it} + \gamma_{it} + \varepsilon_{it}$$
(1)

In equation (1), MEF_{it} refers to the dependent variable capturing management earnings forecast likelihood and properties for firm *i* in year *t*.¹⁸ *Republican* refers to the various proxies that capture CEOs' Republican political ideology; we use *Rep_Dum* or *Rep_Index* in our main analysis and *Rep_dum_{cycle}*, *Rep_dum_{tenure}*, or *Rep_index_{cycle}* in robustness tests. γ is a vector of control variables. We include a set of indicator variables to control for year and industry fixed effects in all models.

We control for firm size (Ln(assets)), market-to-book ratio (MB), financial leverage (Leverage), the intensity of investment in research and development expenditure (RD), return on assets (ROA), return volatility (Volatility), analyst following (Ln(Analyst)), and institutional ownership (Instit Own), because prior research shows that these variables influence the likelihood and properties of MEFs (Miller, 2002; Ajinkya et al., 2005; Houston et al., 2019). We also control for litigation risk (*Litigation*) because MEFs that are made in good faith are inversely associated with the likelihood of litigation (Francis et al., 1994; Matsumoto, 2002). We include news type (News) to control for the direction of the change in earnings per share from the prior year (Baginski et al., 2002). We control for firms' engagement in equity issues (Equity Issue) and in mergers and acquisitions (Acquisition) because firms may provide biased disclosures to reduce information asymmetry when undergoing significant events such as new issue offerings or mergers and acquisitions (Hribar and Yang, 2016). Finally, we control for product market competition (Industry Conc) because firms in highly competitive industries may issue less optimistic forecasts, specifically when investors have difficulty identifying the forecast bias (Rogers and Stocken, 2005). Appendix A provides detailed definitions of the above control variables.

¹⁸ We use an OLS regression model for continuous earnings forecasts measures and a logit regression model for binary earnings forecasts measures.

4.2. CEO political ideology and the likelihood and frequency of issuing MEFs

In this section, we test our first conjecture that firms with Republican CEOs are more likely to issue earnings forecasts, and that once they decide to issue, such firms tend to issue their forecasts more frequently, compared to those with non-Republican CEOs. Table 3 presents the results for the association of CEO political ideology with the likelihood of issuing MEFs and the properties of MEFs.

[Insert Table 3]

In Table 3, models (1) and (2) present results of logistic regression models in which the dependent variable is the likelihood of issuing MEFs (Issue), and models (3) and (4) present results of OLS regression models in which the dependent variable is the frequency of MEFs (Frequency). CEO Republican ideology is measured using Rep Dum in models (1) and (3), and Rep Index in models (2) and (4). In the models for *Issue* (models (1) and (2)), we find that the coefficients on both Rep Dum and Rep Index are positive and highly significant at less than the 1% level. This finding is consistent with our expectation that Republic CEOs are more likely to issue MEFs than non-Republican CEOs. Specifically, in model (1), the coefficient estimate on Rep Dum (0.128; tvalue = 3.86) indicates that Republican CEOs are nearly 13 percent more likely to issue earnings forecasts in any given year compared to non-Republican CEOs. In the models for *Frequency* (model (3) and (4)), we also find that the coefficients on both measures of the Republican political ideology are positive and highly significant at less than the 1% level. The finding is consistent with the view that Republican CEOs tend to issue MEF more frequently than non-Republican CEOs. Specifically, in model (3), the coefficient estimate on *Rep Dum* (0.165; t-value = 5.26) suggests that Republican CEOs, on average, issue 16.5 percent more forecasts compared to non-Republican CEOs. We find similar results in models (2) and (4), where CEO Republican ideology is proxied by *Rep Index*.

The coefficient estimates on control variables, i.e., other determinants of the likelihood and frequency of issuing MEFs, are largely consistent with the findings of prior studies. For instance, we find a positive association between firm size (Ln(assets)) and the frequency of MEFs (e.g., Ajinkya et al. (2005) and Houston et al., 2019). The negative coefficient on RD indicates that R&Dintensive firms make fewer MEFs. Further, we report a positive coefficient on ROA, which lends support to the view that firms with excellent performance are more likely to issue MEFs and tend to disclose them more frequently (e.g., Miller, 2002). Our results also show a positive association between MEFs and the number of analysts (Ajinkya et al., 2005); and a negative association between MEFs and Volatility (Houston et al., 2019). Finally, due to their large holdings, institutional investors demand that firms release more information (Ajinkya et al., 2005). Similarly, Bird and Karolyi (2016) find that institutional ownership significantly improves firm disclosure policy. Consistent with this finding, we find that *Instit Own* is positively associated with both the likelihood and frequency of issuing MEFs, and the association is highly significant. In short, our results reported in columns (1) to (4) in Table 3, taken together, provide strong and reliable evidence that firms run by Republican CEOs tend to disclose more forecasts than non-Publican CEOs even after controlling for other known determinants of the likelihood and frequency of MEFs.

4.3. CEO political ideology and forecast range and horizon

In this section, we test the effect of CEO political ideology on the likelihood of issuing a range forecast on the one hand and its impact on the MEF horizon on the other hand. We conjecture that, because of their preference for avoiding negative market reactions and litigation if they miss

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their own forecasts, Republican CEOs might prefer to issue range over point estimates. Further, because of their stronger preference for an environment of low information asymmetry, Republican CEOs might issue forecasts with a longer horizon. The results of these tests are reported in models (5)–(8) in Table 3.

In Table 3, the dependent variable is *Range* in the logistic regression models (5) and (6), and is *Ln(Horizon)* in the OLS regression models (7) and (8). CEO Republican ideology is measured using *Rep_Dum* in models (5) and (7), and *Rep_Index* in models (6) and (8). Our results show that CEO Republican ideology is positively associated with both forecast range and horizon. Specifically, the coefficient estimate on *Rep_Dum* in model (5) is 0.127, indicating that firms run by Republican CEOs, on average, issue 12.7 percent more range estimates than firms with non-Republican CEOs. Similarly, the average horizon of forecasts made by Republican CEOs is around 11 percent longer than the horizon of forecasts made by other CEOs. Put differently, given that the average forecast horizon in our sample is 68 days, the horizon of forecasts made by Republican CEOs is around 7–8 days longer than those made by other CEOs. The results are qualitatively the same when *Rep_Index* is used as an alternative measure of CEO Republican ideology.

Our results also show that larger firms issue fewer range forecasts and have a longer forecast horizon than smaller firms. Further, the forecast horizon is longer, and the likelihood of issuing range forecasts is greater for firms that are more levered, more profitable, followed by more analysts, have more institutional ownership, and for those that experience an acquisition during the year. In contrast, the forecast horizon is shorter, and the likelihood of issuing range forecasts is smaller for firms that have higher R&D intensity, higher volatility, have a positive change in EPS (*News*), and for those that experience an equity issuance during the year. These results are, in general, consistent with the findings of prior research on determinants of forecast

horizons and forecast range (Baik et al., 2011; Hribar and Yang, 2016; Houston et al., 2019). In short, the results presented in columns (5) to (8) in Table 3 show significantly positive associations of Republican ideology with *Range* and *Horizon*, suggesting that the conservative ideologies of Republican CEOs influence the properties of their MEFs and its influence is incremental over other known determinants of MEFs.

4.4. CEO political ideology and the credibility of earnings forecasts

To obtain further insight into the role of CEO political ideology in their disclosure behaviors, we test the association between CEO political ideology and several other properties of MEFs. This group of forecast properties could together be seen as a proxy for the credibility of MEFs. We conjecture that Republican CEOs, who have a higher preference for loss and ambiguity avoidance than Democrat CEOs, are more likely to avoid negative earnings surprises, reduce information asymmetry, and reduce the risk of litigation by issuing more accurate and less biased forecasts, compared to non-Republican CEOs. To formally test this conjecture, we first examine whether and how CEO Republican ideology is associated with the likelihood of issuing forecasts with optimistic bias (*OptBias*), the likelihood of actual earnings missing an earnings target (*Forecast_Miss*), and forecast accuracy (*Accuracy*). The results of these tests are reported in Table 4.

[Insert Table 4]

Table 4 reports results of logistic regression models in which the dependent variable is *OptBias* in models (1)–(2), and *Forecast_Miss* in models (3)–(4). Models (5) and (6) report results of OLS regression models at which the dependent variable is *Accuracy*. CEO Republican ideology is measured by *Rep_Dum* in models (1), (3), and (5) and by *Rep_Index* in models (2), (4), and (6).

We find the associations of CEO Republican ideology with *OptBias* and *Forecast_Miss* are negative and highly significant at less than the 1% level, whereas its association with *Accuracy* is positive and highly significant. This finding suggests that Republican CEOs are less likely to issue optimistically biased forecasts and less likely to miss their own earnings forecasts. Specifically, on average, firms run by Republican CEOs are 14.3 percent less likely to issue optimistically biased forecasts, and 14.2 percent less likely to miss forecasts in a given year as compared to other CEOs. Stated another way, given conservatives' preference for an environment of low information asymmetry and for avoiding potential litigation, Republican CEOs are more likely to avoid issuing overly optimistic forecasts and to have their forward-looking voluntary forecasts meet actual realizations, compared to non-Republican CEOs.

Moreover, in model (5) of Table 4, where the dependent variable is *Accuracy*, the coefficient estimate on *Rep_Dum* is positive and highly significant at less than the 1% level. Thus, in addition to issuing less optimistic forecasts and missing forecasts less frequently, Republican CEOs tend to issue more accurate forecasts. Specifically, forecasts made by Republican CEOs are, on average, 8.7 percent more accurate than those made by CEOs with other political ideologies.

The coefficients on control variables are also consistent with prior studies. For example, firm size, market-to-book ratio, return on assets, analyst following, institutional ownership, litigation environment, earnings news type, and industry competition are negatively associated with forecast bias and missing forecasts, captured by *OptBias* and *Forecast_Miss*, respectively, and positively associated with forecast accuracy, captured by *Accuracy*. In contrast, leverage, R&D intensity, volatility, and equity issuance are positively associated with forecast bias and missing forecasts, and negatively associated with forecast accuracy.

Next, we examine the association between CEO political ideology and earnings forecasts news type. Specifically, we differentiate between bad and good news forecasts and examine how CEO political ideology affects the issuance of each type. The results of these tests are presented in columns (1)–(4) in Table 5. The key independent variable is *Rep_Dum* in models (1) and (3), and is *Rep_Index* in models (2) and (4). Our findings reveal that Republican CEOs are more likely to issue bad news forecasts compared to non-Republican CEOs. Specifically, on average, firms with Republican CEOs have around 13 percent more bad news forecasts than those with non-Republican CEOs. In contrast, the coefficient estimates of *Rep_Dum* on good news forecasts (models (3) and (4)) are positive but insignificant, indicating that the political ideology does not play a significant role in CEOs' decision to issue good news forecasts.¹⁹

[Insert Table 5]

These findings are consistent with the notion that Republican CEOs, who are more conservative, voluntarily disclose bad news forecasts to mitigate information asymmetry and avoid future litigation, for example, in case of large stock price declines (Skinner, 1994). Moreover, we find that the coefficient on litigation is positively and significantly associated with bad news forecasts. We also find that firm size, return on assets, analyst following, institutional ownership, earnings news type, involvement in an acquisition, and industry competition are positively associated with bad news forecasts, whereas research and development intensity and volatility are negatively associated.²⁰

¹⁹ We test the estimates of positive versus negative news forecasts using the seemingly unrelated bivariate probit regression (SUR) model and find similar results. These results are available upon request.

²⁰ In contrast, we find that Democrat CEOs are less likely to issue forecasts, have lower forecast frequency, are less likely to issue forecasts in a timely fashion, have lower forecast accuracy, have more optimistically biased forecasts, and are more likely to miss their own forecasts. Table A16 in the Internet Appendix provides tabulated results using *Dem Dummy* and *Dem Index*.

4.5. CEO political ideology and earnings surprises

Skinner and Sloan (2002) show that the market response to negative earnings surprises is much stronger than the market response to positive earnings surprises. Corporate managers thus take actions to avoid negative earnings surprises to avoid their negative impact on stock prices. Matsumoto (2002) argues that managers voluntarily disclose bad news forecasts, or forecasts that are lower than expected, in order to guide analysts' earnings expectations downward, and thus to avoid missing expectations at the earnings announcement date. Although negative earnings surprises are detrimental to firms with both Republican and Democrat CEOs, Republican CEOs are expected to be more sensitive to such incidents because of their higher preference for the avoidance of loss and ambiguity. Consistent with this expectation, our earlier results show that Republican CEOs are more likely to issue bad news forecasts. We, therefore, conjecture that firms run by Republican CEOs would be more (less) likely to experience positive (negative) earnings surprises, compared to firms run by non-Republican CEOs. Results for testing of this conjecture are reported in columns (5)–(10) in Table 5.

In Table 5, the dependent variable in logistic regression is *Positive_Surprise* in models (5)– (6), *Negative_Surprise* in models (7)–(8), and *Neutral_Surprise* in models (9)–(10). The key independent variable is *Rep_Dum* in models (5), (7), and (9), and is *Rep_Index* in models (6), (8), and (10). As shown in Table 5, we find that firms run by Republican CEOs are more likely to experience positive earnings surprises and less likely to experience negative earnings surprises than other firms, which is consistent with our expectation. Specifically, in model (5), where *Positive Surprise* is the dependent variable, the coefficient estimate on *Rep_Dum* is 0.058 (t-value = 2.07), suggesting that firms with Republican CEOs are 5.8 percent more likely to have positive earnings surprises than firms with non-Republican CEOs. In sharp contrast, in model (7), where *Negative Surprise* is the dependent variable, the same coefficient estimate is -0.090 (t-value = - 2.97), indicating that firms with Republican CEOs are 9 percent less likely to have negative earnings surprises than firms with non-Republican CEOs. However, as shown in models (9) and (10), we find that the coefficients on *Rep_Dum* and *Rep-Index* are both insignificant, suggesting that CEO Republican ideology plays no significant role in determining the likelihood of having neutral earnings surprises.

We also find that research and development intensity, return on assets, analysts following, institutional ownership, and earnings news type are positively associated with positive earnings surprises and negatively associated with negative earnings surprises. In contrast, firm size, leverage, and volatility are negatively associated with positive earnings surprises and positively associated with negative earnings surprises.

Collectively, our results presented in columns (5) to (8) in Table 5 lend additional support to our earlier finding that Republican CEOs are more likely than other CEOs to make bad news forecasts. In short, Republican CEOs tend to impose downward pressure on analysts' forecasts, aiming at a higher probability of experiencing positive earnings surprises.

5. CEO political ideology and MEFs: Identification and endogeneity issues

Our baseline findings show strong associations between CEO political ideology and both the likelihood of issuing MEFs and their properties. However, one can argue that our findings could be driven by endogenous firm-CEO matching. For instance, firms with higher disclosure quality may tend to appoint Republican CEOs, and/or Republican CEOs might tend to move to firms that have a superior disclosure quality environment. Similarly, directors and top executives may prefer to hire a CEO who shares their political affiliation, and/or a CEO might prefer to work in a company whose directors and top executives share her/his political affiliation. For instance, Wintoki and Xi (2019) show that fund managers prefer to allocate assets to firms managed by executives and directors with whom they share a similar political affiliation. More recently, Twitter CEO and co-founder Jack Dorsey, who exclusively donates to Democrats,²¹ was subject to a severe threat of losing his position after the well-known Republican activist investor Elliot Management Corporation purchase a sizable stake in Twitter.²² However, Khanna et al. (2015) argue that such a connection between the CEO and other top executives increases the risk of corporate fraud and reduces the likelihood of CEO dismissal upon the discovery of such fraud. Lee et al. (2014) show that alignment in political orientation between the CEO and independent directors is associated with lower firm valuations, lower operating profitability, and increased internal agency conflicts such as lower turnover for poorly performing CEOs and lower pay-performance sensitivity.

To alleviate concerns about such biases arising from endogenous firm-CEO matching, we conduct multiple causality tests using the propensity score matched sample and performing a difference-in-differences (DID) analysis.²³ To further address concerns about correlated omitted variables, we conduct additional tests that control for CEO characteristics, incentives, and overconfidence. Finally, to address possible error-in-measurement issues with our baseline proxy for CEO political ideology, we use several alternative measures of political ideology following Hutton et al. (2014) and Elnahas and Kim (2017).

²¹ See more details at https://nypost.com/2018/08/04/how-twitter-is-fueling-the-democratic-agenda// (accessed 7 May 2020).

²² See more details at https://www.bloomberg.com/news/articles/2020-02-29/singer-s-elliott-is-said-to-seek-to-replace-twitter-ceo-dorsey (accessed 7 May 2020).

²³ In our previous tests, we controlled for various firm and CEO characteristics, and year and industry fixed effects. We also perform robustness tests controlling for state fixed effects, run subsample tests excluding CEO turnover years and excluding the first 3 years of the CEO's tenure, and perform change-on-change regressions. These results are available in the Internet Appendix.

5.1. Propensity score matching

In this section, we use the propensity score matching (PSM) technique to construct a treatment group of firms with Republican CEOs and a control group of firms with non-Republican CEOs. Specifically, the *Treatment* group is identified using Rep_dum_{cycle} , which is an indicator variable that equals one if all donations made by a CEO in an election cycle are directed to the Republican Party (with none to the Democratic Party), and zero otherwise. As the first step of the PSM, we estimate a logistic regression of Rep_dum_{cycle} on multiple firm characteristics as in Panel A of Table 6. Using the estimated coefficients from the logistic model, we then compute the propensity scores (i.e., the predicted likelihood) of $Rep_dum_{cycle} = 1$ for all firms in our sample. We then match each treated firm with a control firm that has the nearest neighbor propensity score. As the second step of the PSM, we re-estimate all regressions in Tables 3, 4, and 5 using the PSM-screened sample. We report the PSM results using Rep_dum_{cycle} as the key variable of our interest in Table 6.

[Insert Table 6]

Panel A in Table 6 presents results for the diagnostic differences in means of firm characteristics between the *Treatment* and *Control* groups. Reported t-statistics show no statistically significant differences in firm characteristics between the *Treatment* and *Control* groups. Panels B and C in Table 8 present the results for regressions of various properties of MEFs on CEO political ideology, using the propensity score-matched firm-year observations. Our findings indicate that Rep_dum_{cycle} is positively associated with forecast issue, frequency, range, horizon, accuracy, bad news forecast, and positive earnings surprises, while it is negatively associated with the missing forecast, bias, and negative earnings surprises. Overall, even though our matched sample size (N = 9,578) is substantially smaller than our main sample (N = 33,951),

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the effect of CEO political ideology on the likelihood of issuing MEFs and various MEF properties is qualitatively similar to that reported in our baseline models.²⁴

5.2. Management earnings forecasts around CEO Turnover: Difference-in-differences tests

To better establish the causal relation between CEO political ideology and MEFs, we exploit CEO turnovers (from Republican CEOs to non-Republican CEOs or vice versa) as a setting in which to apply the DID analysis of how changes in CEO political ideology influence various properties of MEFs. To this end, we construct a new indicator variable, *Rep-Leaving*, that equals one if a firm replaces a Republican CEO with a non-Republican CEO, and zero otherwise. Republican CEOs are identified using *Rep dum_{Only}*, which is an indicator variable that equals one if all donations made by a CEO in an election cycle are directed to the Republican Party, and zero otherwise. We also create an indicator variable, After, that equals one for post-turnover years, and zero for pre-turnover years. The interaction term, *Rep-Leaving*×*After*, captures the DID effect of replacing a Republican CEO with a non-Republican CEO on the likelihood of issuing MEFs and various properties of MEFs once issued. Based on our baseline results, we predict reductions in the likelihood of issuing MEFs, forecast frequency, and the quality and credibility of MEFs in the post period of CEO turnovers (After = 1). To avoid the impact of other confounding effects, we use firm-year observations for the window of (-3, +3) years around each CEO turnover event. Further, we restrict our test to turnover events where a long-term incumbent CEO is replaced by a long-term new CEO, where long-term CEOs are those who hold their position for at least three years. Table 7 reports the results of these DID tests.

[Insert Table 7]

²⁴ Our results also remain qualitatively similar if we reconstruct the treatment and control groups based on our alternative measures of CEO political ideology. These results are available in the Internet Appendix.

Consistent with our baseline results, the coefficient on *Rep-Leaving*×*After* is negative and statistically significant in models examining MFE *Issue*, *Frequency*, *Range*, *Horizon*, *Accuracy*, and *Bad_News*, whereas the same coefficient is positive and statistically significant in models examining *Forecast_Miss*, *Bias*, and *Negative_Surprise*.²⁵ Overall, these results lend strong support to the view that our main results are unlikely to be driven by potential endogeneity, particularly reverse causality. Replacing a Republican CEO with a non-Republican CEO leads to lowering the likelihood of issuing MEF and the issue frequency, and deteriorating the quality and credibility of MEFs.

5.3. Controlling for CEO characteristics, incentives, and overconfidence

Our proxies for CEO Republican ideology are constructed from each CEO's individual donations data and hence are likely to capture CEO political orientation reasonably well. However, if these proxies are mere reflections of, and/or highly correlated with, other CEO characteristics that our baseline models do not control for, then our baseline results could suffer from potential problems that may arise from correlated omitted variables. To alleviate concerns about such problems, we control for a wide range of CEO characteristics. Specifically, we control for Ln(Tenure) and Duality because forecasting accuracy and earnings announcement tone are shown to be positively associated with the managers' experience and CEO duality (Feng et al., 2009). We also control for CEO gender because prior research shows that female CEOs are less likely to engage in opportunistic and fraudulent behavior (e.g., Ali and Hirshleifer 2017). We also control for Ln(Age) because prior studies find a negative relation between CEO age and bad news hoarding

²⁵ Our findings are qualitatively similar if we restrict our sample to -2, +2 years around CEO turnover events. We also find statistically significant results for changes in earnings forecasts associated with changes in CEO political ideology due to CEO turnover where $\triangle REP_{CEO} = 1$ if a Republican CEO (Rep_dum_{Only}) replaces a Democratic CEO (Dem_dum_{Only}), 0 if the political ideology is similar after a CEO turnover, and -1 if a Democratic CEO replaces a Republican CEO. These results are available in the Internet Appendix.

(Andreou et al., 2017) and a positive relation between CEO age and financial reporting quality (Huang et al., 2012). We include CEO pay-performance sensitivity (Ln(Delta)) and CEO risk-taking incentive (Ln(Vega)) because prior research finds a positive relation between CEO equity compensation and MEFs (Baginski et al., 2018). Similarly, managers may opportunistically provide voluntary disclosure to maximize their stock option compensation (Cheng et al., 2013). We also control for measures of CEO overconfidence, such as *Holder67* and *Net_buyer* following Hribar and Yang (2016), and CEO ownership (CEO_Own) following Malmendier and Tate (2005). The results of these tests are reported in Table 8.

[Insert Table 8]

Panel A in Table 8 reports results for the association of CEO political ideology with *Issue*, *Frequency*, *Range*, *Ln*(*Horizon*), *OptBias*, and *Forecast_Miss*, and Panel B reports results for its association with *Accuracy*, *Bad_News*, *Good_News*, *Positive_Surprise*, *Negative_Surprise*, and *Neutral_Surprise*. Even after controlling for the aforementioned firm and CEO characteristics, the results are qualitatively identical with those of the baseline models. Specifically, we find a positive association between measures of CEO Republican ideology and such forecast-related variables as the likelihood of forecast issuance, forecast frequency, range, horizon, accuracy, and bad news. We further find a negative association between measures of CEO Republican ideology and such forecasts, and negative earnings surprises. Consistent with prior studies, we find that *Vega* is positively associated with the likelihood of forecast issuance, forecast issuance, forecast frequency, range, horizon, accuracy, and bad news, while it is negatively associated with the forecast bias and missing forecast (Baginski et al., 2018). We also find that CEO overconfidence is positively associated with the likelihood of forecast frequency. In addition, we find that female CEOs tend

to issue more guidance, issue long horizon guidance, and are associated with less earnings surprises, which is consistent with the findings that female CEOs are likely to be more conservative in disclosure.

5.4. Alternative measures of CEO political ideology

Our baseline models use *Rep_Dum* (Bhandari et al., 2018) and *Rep_Index* (Hong and Kostovetsky, 2012) as proxies for CEO Republican ideology. The construction of these individual proxies may represent another source of potential endogeneity in our baseline results due to measurement error. In an effort to mitigate potential bias and noise in the baseline measures of CEO political ideology, we employ three alternative measures of CEO Republican ideology, namely *Rep_indexcycle*, *Rep_dumcycle* (Hutton et al., 2014), and *Rep_dumtenure* (Elnahas and Kim, 2017). Table 9 reports the test results using these alternative measures.

[Insert Table 9]

The results in Table 9 are, overall, in line with our baseline results, and alleviate the concern that our main findings are biased owing to measurement errors in individual proxies for Republican ideology. Specifically, we find that Republican CEOs are, on average, 8 to 12 percent more likely to issue forecasts compared to non-Republican CEOs (depending on the alternative Republican ideology measure used). Further, on average, Republican CEOs have a 9 to 11 percent greater forecasting frequency than non-Republican CEOs. Similarly, using these alternative proxies for political ideology, Republican CEOs consistently have a higher likelihood of issuing range forecasts, longer forecast horizon, higher forecast accuracy, more likelihood of issuing bad news forecasts, higher likelihood of having earnings surprises, less forecast bias, and a lower likelihood of missing earnings forecasts.

5.5. Republican CEOs: The authoritarian effect versus the precautionary effects

Our results so far are consistent with the *precautionary* effect for Republican CEOs, and are inconsistent with the *authoritarian* effect. In order to further establish the precautionary effect as an explanation for our results, we conduct cross-sectional tests using subsamples of firms with high and low institutional ownership and with high and low litigation risk. If Republican CEOs indeed adopt high-quality MEF policies as a precaution to avoid litigation and career penalties, then our results are expected to be stronger for firms with stronger institutional monitoring and higher litigation risk. We report the results of our cross-sectional tests based on institutional ownership in Table 10.

[Insert Table 10]

In Table 10, Panels A1 and A2 report results for the subsample of firms with high (abovemedian) institutional ownership, and Panels B1 and B2 report results for firms with low (belowmedian) institutional ownership. The impact of CEO Republican ideology on the likelihood of issuing and characteristics of MEFs is much stronger in the high institutional ownership subsample. For example, using *Rep_Dum*, in firms with high institutional ownership, the likelihood of issuing MEFs is 17.7 percent higher for Republican CEOs than for non-Republican CEOs (Column (1); Panel A1), whereas in firms with low institutional ownership, the likelihood is only 9.7 percent higher for Republican CEOs than for other CEOs (Column (1); Panel B1). Similarly, in firms with high institutional ownership, Republican CEOs are 20.2 percent more likely to issue range forecasts than non-Republican CEOs (Column (5); Panel A1), whereas in firms with low institutional ownership, they are only 1.9 percent more likely to do so (Column (5); Panel B1). These results are consistent across all other variables that capture different characteristics of MEFs.

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In Table 11, we report the results of our cross-sectional tests based on litigation risk. Panels A1 and A2 report results for the subsample of firms with high litigation risk, and Panels B1 and B2 report the same for the subsample of firms with low litigation risk. The impact of CEO Republican ideology on the likelihood and characteristics of MEFs is again much stronger in the high litigation risk subsample. For example, using *Rep_Dum*, in firms with high litigation risk, the *Accuracy* of MEFs of Republican CEOs is 20.6 percent higher than non-Republican CEOs (Column (1): Panel A2), whereas in firms with low litigation risk, it is only 5.4 percent higher (Column (1); Panel B2). These results are also consistent across most of the other characteristics of MEFs. In general, these cross-sectional results lend strong support to the *precautionary effect* explanation. Republican CEOs favor more frequent and higher-quality forecasts when the likelihood of disciplinary action is elevated due to strong institutional monitoring or high litigation risk.

[Insert Table 11]

7. Conclusion

The main premise of this paper is that a CEO's political ideology can translate into her/his decisions related to voluntary disclosure. Specifically, Republican CEOs, who are often described as more conservative, might use voluntary disclosure to reduce information asymmetry, the likelihood of negative earnings surprises, and the risk of litigation.

Our results, using CEOs' political contributions data for the period 1993–2016, show that firms run by Republican CEOs are more likely to issue forecasts and have higher forecast frequency than other firms. Consistent with the conservative characteristics of Republican CEOs, we find that they are more likely to issue range forecasts, have less forecasting bias, and less likely

to subsequently miss their own forecast. Republican CEOs also issue forecasts in a timelier fashion and with higher accuracy than other CEOs. Further analyses reveal that Republican CEOs tend to issue negative news, leading to more positive earnings surprises. Our results are robust to the use of several alternative measures of CEO political ideology to address potential error-inmeasurement issues. Further, we obtain consistent results when using multiple subsample analyses.

In short, our results provide strong and reliable evidence that CEO political ideology does affect corporate policy choices, specifically relating to the voluntary (and thus discretionary) disclosure of forward-looking information. Given the scarcity of empirical evidence on the role of CEO political ideology in shaping firms' accounting and auditing policies, we recommend further research in this direction. In particular, a fruitful area for future research would be to investigate whether and how CEO political ideology influences the quality of financial reporting, including its qualitative nature such as readability, tone, and other linguistic quality of narrative disclosure.

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Variable	Definition
CEO political ideo	
Rep_Dum	An indicator variable that equals one if a CEO donated more to the Republican Party than to the Democratic Party during their tenure [Bhandari et al., 2018].
Rep_Index	The percentage of a CEO's support for the Republican Party, calculated as the number of cycles in which a CEO donates exclusively to the Republican Party divided by their number of donation cycles in the sample period [Hong and Kostovetsky, 2012].
CEO political ideo	
Dem_Dum	An indicator variable that equals one if a CEO donated more to the Democratic Party than to the Republican Party during their tenure, and zero otherwise.
Dem_Index	The percentage of a CEO's support for the Democratic Party, calculated as the number of cycles in which a CEO donates exclusively to the Democratic Party divided by their number of donation cycles in the sample period.
Rep_dum _{cycle}	An indicator variable that equals one if all donations by a CEO in an election cycle are directed to the Republican Party, and zero otherwise [Hutton et al., 2014].
Rep_dum_{tenure}	An indicator variable that equals one if all donations by a CEO during their tenure are directed to the Republican Party, and zero otherwise [Elnahas and Kim, 2017].
<i>Rep_index</i> _{cycle}	An index calculated as total donations to the Republican Party minus total donations to the Democratic Party divided by total donations to both parties in each election cycle. This index ranges between -1 (strong Democrat) and 1 (strong Republican) [Hutton et al., 2014].
Rep_dum _{Only}	An indicator variable that equals one if all donations by a CEO in an election cycle are directed to the Republican Party only (and none to the Democratic or other parties), and zero otherwise.
Voluntary disclosu	ire
Issue	An indicator variable that equals one if a firm makes annual earnings forecasts in a fiscal year, and zero otherwise.
Frequency	The total number of annual earnings forecasts made by a firm in a fiscal year.
Ln(Horizon)	The natural logarithm of one plus the average horizon of annual earnings forecasts made by a firm in a fiscal year. For each forecast, the horizon is defined as the number of calendar days between the forecast announcement date and the corresponding period end date. If a firm makes no forecasts in a fiscal year, we assign an average horizon value of zero.
Range	An indicator variable denoting that a firm issues range estimates. For each forecast, we first assign a value of one for range estimates, and zero otherwise. These individual forecast values are then averaged for each firm-year. <i>Range</i> is then defined as an indicator variable that equals one if the average range is greater than 0.5, and zero otherwise.
Forecast_Miss	An indicator variable that equals one if a firm's earnings fail to meet forecast at least once in a year, and zero otherwise. For open-ended and point forecasts, <i>Forecast_Miss</i> equals one if the actual earnings are less than the MEF, and zero otherwise. For range estimates, the variable equals one if the actual earnings are less than the lower bound of the range forecast, and zero otherwise.
<i>OptBias</i>	An indicator variable that equals one if the average <i>Bias</i> in a year is positive, and zero otherwise, where for each estimate, <i>Bias</i> is the difference between MEFs and actual earnings scaled by the stock price at the end of the month prior to the forecast.

Appendix A. Variable definitions.

Appendix A. Variable definitions. Cont'd

Appendix A. Variat Accuracy	ble definitions. Cont'd The average forecast accuracy for all annual earnings forecasts made by a firm in
Accuracy	a fiscal year. For each estimate, we first calculate the absolute difference between MEFs and actual earnings scaled by the stock price at the end of the month prior to the forecast. Next, we identify forecast accuracy as the quintile ranking of the scaled forecast difference, where one is assigned for the top quintile (largest error), and five is assigned for the bottom quintile (lowest error). The value zero is
	assigned if no forecasts are made.
Bad_News	An indicator variable that equals one if forecast news is negative, and zero otherwise. Forecast news is the difference between the MEF and the most recent mean analyst estimate scaled by the stock price one trading day prior to the MEF release date.
Good_News	An indicator variable that equals one if forecast news is non-negative, and zero otherwise. Forecast news is the difference between the MEF and the most recent mean analyst estimate scaled by the stock price one trading day prior to the MEF release date.
Positive_Surprise	An indicator variable that equals one if an earnings surprise is greater than 0.0001, and zero otherwise. The earnings surprise is calculated as the difference between the actual earnings and the mean analyst estimate scaled by the stock price three trading days prior to the earnings announcement.
Negative_Surprise	An indicator variable that equals one if an earnings surprise is less than -0.0001, and zero otherwise. The earnings surprise is calculated as the difference between the actual earnings and the mean analyst estimate scaled by the stock price three trading days prior to the earnings announcement.
Neutral_Surprise	An indicator variable that equals one if an earnings surprise is between 0.0001 and -0.0001, and zero otherwise. The earnings surprise is calculated as the difference between the actual earnings and the mean analyst estimate scaled by the stock price three trading days prior to the earnings announcement.
Firm Characteristics	
Ln(assets)	The natural logarithm of total assets. [at]
MB	The ratio of market-to-book value of equity. [(prcc_f*csho) / ceq]
Leverage	The ratio of total debt to market value of total assets. [(Dltt+Dlc) / (at-ceq+csho*prcc f)]
RD ROA	Expenditure on research and development scaled by total assets. [xrd/at] Return on assets, measured as income before extraordinary items scaled by total
Volatility	assets. [ib/at] The standard deviation of daily stock returns (CRSP variable ret) of a firm over the last fiscal year.
Ln(Analyst)	The natural logarithm of the number of analysts following a firm.
Instit Own	The percentage of shares owned by institutional investors.
Litigation	An indicator variable that equals one if a firm's SIC code denotes an industry subject to increased litigation (2833-2836, 3570-3577, 3600-3674, and 7370-7374), and zero otherwise.
News	An indicator variable that equals one if the current period EPS is greater than or equal to the previous-period EPS, and zero otherwise.
Equity_Issue	An indicator variable that equals one if a firm issued shares in a year, and zero otherwise.
Acquisition	An indicator variable that equals one if a firm's annual acquisition- or merger- related costs exceed five percent of net income (or loss) in the year, and zero otherwise. [aqc/ni]

Industry_Conc	A firm's industry concentration, measured as the sum of sales of the top five firms
	in its two-digit SIC code scaled by total sales of all firms in its two-digit SIC code
	in the year. $\left[\sum_{i=1}^{5} Sale_{i,j} / \sum_{i=1}^{n} Sale_{i,j}\right]$

CEO Characteristi	cs
Ln(Tenure)	The natural logarithm of CEO tenure, where tenure is defined as the length of a
	CEO's tenure with their current firm (measured as the fiscal year minus the year
	they joined as CEO).
Ln(Age)	The natural logarithm of the age of a CEO in the year in which a MEF was released.
Duality	An indicator variable that equals one if a CEO is also the chairperson of the firm's board, and zero otherwise.
CEO Gender	CEO Gender equals 1 if a CEO is female, 0 otherwise.
Ln(Delta)	The natural logarithm of the expected dollar change in CEO wealth for a 1 percent change in stock price, computed as in Core and Guay (2002).
Ln(Vega)	The natural logarithm of the expected dollar change in CEO wealth for a 1 percent change in stock return volatility, computed as in Guay (1999).
CEO_Own	The percentage of outstanding shares owned by a CEO. [SHROWN_EXCL_OPTS / (CSHO [*] 1000)]
Holder67	An indicator variable that equals one if a CEO holds vested options with average moneyness greater than 67 percent, and zero otherwise starting in the first year when a CEO displays this behavior. Option moneyness is calculated as follows: First, we calculate the realizable value per option as the total realizable value of the exercisable options divided by the number of exercisable options [Value_Per_option = (OPT_UNEX_EXER_EST_VAL / OPT_UNEX_EXER_NUM)]. Second, we compute the estimate of the average exercise price of the options by subtracting the per-option realizable value from the stock price at the fiscal year-end [avg_exercise_price = (prccf - Value_Per_option)]. Finally, the average percent moneyness of an option equals the per-option realizable value divided by the estimated average exercise price [avg_pctg_moneyness_opt = (Value_Per_option / avg_exercise_price)]. (Malmendier and Tate, 2005; Campbell et al., 2011; Hirshleifer et al., 2012.)
Net_buyer	An indicator variable that equals one if the number of years at which a CEO is a net-buyer is higher than those at which she/he is a net seller. <i>Net_buyer</i> is calculated as follows: first, we compute the net stock purchases by a CEO as purchases minus sales, both in units of shares [net_purchase = ($SHROWN_EXCL_OPTS_t - SHROWN_EXCL_OPTS_{t-1}$], then we calculate the number of years at which a CEO has bought more shares than he/she sold. [Malmendier and Tate, 2005; Campbell et al. 2011]

Appendix A. Variable definitions. Cont'd

Table 1. Summary Statistics.

This table reports descriptive statistics for measures of CEO political ideology, voluntary disclosure, and control variables for our sample covering the period 1993-2016. *Rep_Dum* is an indicator variable that equals one if a CEO donated more to the Republican Party than to the Democratic Party during her/his tenure. *Rep_Index* is the percentage of a CEO's support for the Republican Party calculated as the number of cycles in which a CEO donates exclusively to the Republican Party divided by her/his number of donation cycles in the sample period. All other variables are defined in Appendix A.

variable	No.	Mean	Std Dev	25 th Perc	Median	75 th Perc
CEO political ideology						
Rep_Dum	33,951	0.229	0.420	0.000	0.000	0.000
Rep_Index	33,951	0.169	0.337	0.000	0.000	0.000
Voluntary Disclosure						
Issue	33,951	0.353	0.478	0.000	0.000	1.000
Frequency	33,951	1.550	2.619	0.000	0.000	3.000
Ln(Horizon)	33,951	1.828	2.493	0.000	0.000	5.141
Range	33,951	0.266	0.442	0.000	0.000	1.000
Forecast_Miss	33,951	0.190	0.392	0.000	0.000	0.000
OptBias	33,951	0.143	0.350	0.000	0.000	0.000
Accuracy	33,951	1.035	1.592	0.000	0.000	2.000
Bad_news	33,951	0.278	0.448	0.000	0.000	1.000
Good News	33,951	0.192	0.394	0.000	0.000	0.000
Positive_Surprise	33,951	0.530	0.499	0.000	1.000	1.000
Negative_Surprise	33,951	0.368	0.482	0.000	0.000	1.000
Neutral_Surprise	33,951	0.102	0.302	0.000	0.000	0.000
Firm Characteristics						
Ln(Assets)	33,951	7.181	1.597	6.023	7.047	8.211
MB	33,951	3.235	4.049	1.488	2.352	3.844
Leverage	33,951	0.147	0.144	0.020	0.114	0.225
RD	33,951	0.034	0.059	0.000	0.003	0.044
ROA	33,951	0.037	0.114	0.015	0.052	0.090
Volatility	33,951	0.027	0.014	0.018	0.024	0.034
Ln(Analyst)	33,951	2.116	0.805	1.609	2.197	2.708
Instit_Own	33,951	0.542	0.360	0.150	0.647	0.837
Litigation	33,951	0.240	0.427	0.000	0.000	0.000
News	33,951	0.631	0.482	0.000	1.000	1.000
Equity_Issue	33,951	0.203	0.402	0.000	0.000	0.000
Acquisition	33,951	0.411	0.492	0.000	0.000	1.000
Industry_Conc	33,951	0.475	0.151	0.356	0.447	0.539

Table 2. Pearson correlations

This table reports Pearson correlation coefficients. *Rep_Dum* and is an indicator variable that equals one if a CEO donated more to the Republican Party than to the Democratic Party during her/his tenure. *Rep_Index* is the percentage of a CEO's support for the Republican Party calculated as the number of cycles in which a CEO donates exclusively to the Republican Party divided by her/his number of donation cycles in the sample period. All other variables are defined in Appendix A.

Variables		Α	В	С	D	Е	F	G	Η	Ι	J	Κ	L	М	Ν	0	Р	Q	R	S	Т	U
Rep_Dum	А	1																				
Rep_Index	В	0.86																				
Issue	С	0.03	0.04																			
Frequency	D	0.04	0.05	0.80																		
Ln(Horizon)	Е	0.03	0.04	0.99	0.80																	
Range	F	0.02	0.03	0.85	0.74	0.85																
Forecast_Miss	G			-0.60																		
OptBias	Η	-0.03	-0.04	-0.70	-0.59	-0.69	-0.63	0.85														
Accuracy	Ι	0.04	0.05	0.88	0.73	0.87	0.75	-0.58	-0.69													
Ln(Assets)	J	0.15	0.21	0.20	0.24	0.20	0.17	-0.13	-0.16	0.22												
MB	Κ			0.05																		
Leverage	L			-0.04																		
RD	Μ			-0.08				0.04	0.05	-0.08	-0.27	0.13	-0.30									
ROA	Ν	0.05	0.06	0.13	0.13	0.13	0.11	-0.15	-0.17	0.18	0.14	0.17	-0.25	-0.29								
Volatility	0	-0.08	-0.09	-0.13	-0.17	-0.13	-0.13	0.10	0.12	-0.18	-0.35	-0.06	0.00	0.25	-0.29							
Ln(Analyst)	Р			0.19																		
Instit_Own	Q			0.12																		
Litigation	R			-0.02																		
News	S			0.00								-	-						-0.00			
Equity_Issue	Т																		0.21			
Acquisition	U	0.00	0.00	0.14	0.13	0.13	0.12	-0.07	-0.09	0.14	0.12	-0.03	0.05	-0.07	0.07	-0.13	0.06	0.11	0.00	-0.03	-0.00	
Industry_Conc	V	0.02	0.02	0.04	0.07	0.04	0.05	-0.03	-0.04	0.04	0.06	-0.02	0.05	-0.04	0.02	-0.07	-0.07	0.03	-0.16	-0.00	-0.07	0.02

Table 3. CEO political ideology and the Likelihood, Frequency, Range, and Horizon of MEF. This table presents tests of the association between CEO political ideology and the likelihood and frequency of earnings forecasts as well as the likelihood of issuing range forecasts and forecast horizon. In the Logistic regressions in columns (1) and (2), the dependent variable is *Issue* which is an indicator variable that equals one if a firm makes annual earnings forecasts in a fiscal year, and zero otherwise. In the OLS regressions in columns (3) and (4), the dependent variable is *Frequency*, which is the total number of annual earnings forecasts made by a firm in a fiscal year. In the Logistic regression in columns (5) and (6), the dependent variable is *Range*, which is an indicator variable of range estimates. In the OLS regressions in columns (7) and (8), the dependent variable is *Ln(Horizon)*, which is the natural logarithm of one plus the average horizon of annual earnings forecasts made by a firm in a fiscal year. Measures of CEO political ideology, *Rep_Dum* and *Rep_Index, and all* other independent variables are defined in Appendix A. All models include year and industry fixed effects. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

	Iss	sue		uency		nge	Ln(Ho	orizon)
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Rep_Dum	0.128^{\dagger}		0.165^{\dagger}		0.127^{\dagger}		0.111^{\dagger}	
	(3.86)		(5.26)		(3.62)		(3.86)	
Rep_Index		0.126 [†]		0.117^{\dagger}		0.144^{\dagger}		0.100^{\dagger}
		(3.08)		(3.04)		(3.34)		(2.82)
Ln(assets)	0.018	0.021	0.121^{+}	0.126 [†]	-0.025*	-0.022	0.022^{*}	0.024^{**}
	(1.28)	(1.51)	(9.72)	(10.09)	(-1.71)	(-1.53)	(1.85)	(2.09)
MB	0.002	0.002	0.015^{\dagger}	0.015 [†]	0.000	0.000	0.008**	0.008**
	(0.44)	(0.49)	(4.12)	(4.20)	(0.01)	(0.06)	(2.37)	(2.42)
Leverage	0.450†	0.449†	0.364 [†]	0.362†	0.380†	0.380 [†]	0.246**	0.244**
-	(3.66)	(3.65)	(3.68)	(3.66)	(2.90)	(2.90)	(2.48)	(2.46)
RD	-3.649†	-3.637 [†]	-1.935†	-1.925†	-4.655†	-4.638 [†]	-2.600†	-2.591*
	(-9.74)	(-9.71)	(-7.14)	(-7.10)	(-11.01)	(-10.97)	(-9.62)	(-9.58)
ROA	1.406†	1.413 [†]	0.790 [†]	0.796†	1.242†	1.247	0.778^{+}	0.782 [†]
	(7.97)	(8.01)	(7.26)	(7.31)	(6.48)	(6.51)	(6.51)	(6.54)
Volatility	-23.445†	-23.389†	-17.674†	-17.654†	-23.294†	-23.230†	-18.934†	-18.917
·	(-16.69)	(-16.66)	(-17.64)	(-17.62)	(-15.47)	(-15.44)	(-17.43)	(-17.42)
Ln(Analyst)	0.671	0.671*	0.467†	0.469†	0.546†	0.546†	0.542†	0.543†
	(26.76)	(26.78)	(23.32)	(23.40)	(20.66)	(20.66)	(27.38)	(27.42)
Instit Own	0.251 [†]	0.248 [†]	0.184 [†]	0.180 [†]	0.216†	0.213 [†]	0.185 [†]	0.182 [†]
—	(6.06)	(5.99)	(4.54)	(4.45)	(4.96)	(4.88)	(4.82)	(4.75)
Litigation	0.209†	0.210 [†]	0.289†	0.289 [†]	0.186 [†]	0.189 [†]	0.205†	0.205†
0	(3.67)	(3.69)	(5.48)	(5.46)	(3.16)	(3.21)	(4.11)	(4.12)
News	-0.115†	-0.115†	0.013	0.013	-0.101†	-0.101†	-0.083†	-0.082†
	(-3.82)	(-3.81)	(0.47)	(0.50)	(-3.17)	(-3.17)	(-3.19)	(-3.17)
Equity Issue	-0.094***	-0.095**	-0.025	-0.026	-0.134 [†]	-0.134 [†]	-0.080 [†]	-0.081 [†]
	(-2.46)	(-2.48)	(-0.79)	(-0.82)	(-3.22)	(-3.23)	(-2.58)	(-2.60)
Acquisition	0.353 [†]	0.353 [†]	0.302^{\dagger}	0.302^{\dagger}	0.334 [†]	0.334 [†]	0.321 [†]	0.321 [†]
	(12.27)	(12.28)	(11.17)	(11.17)	(10.93)	(10.95)	(12.49)	(12.49)
Industry Conc	0.704^{\dagger}	0.705 [†]	0.554^{\dagger}	0.549 [†]	0.290*	0.294*	0.552^{\dagger}	0.550 [†]
· _	(4.99)	(4.99)	(4.56)	(4.51)	(1.92)	(1.94)	(4.72)	(4.71)
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951
Pseudo / Adj. R^2	0.257	0.257	0.280	0.279	0.253	0.253	0.269	0.268

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Table 4. CEO Political Ideology and MF Bias, Miss, and Accuracy

This table presents tests of the association between CEO political ideology and management earnings forecast's optimism, miss incidences, and accuracy. In the Logistic regressions in columns (1) and (2), the dependent variable is *OptBias*, which is an indicator variable that equals one if the average bias in a year is positive, and 0 otherwise. In the Logistic regressions in columns (3) and (4), the dependent variable is *Forecast_Miss*, which is an indicator variable that equals one if a firm misses at least one forecast in a year, and zero otherwise. In the OLS regressions in columns (5) and (6), the dependent variable is *Accuracy*, which is the average Forecast accuracy for all annual earnings forecasts made by a firm in a fiscal year. Measures of CEO political ideology, *Rep_Dum*, and *Rep_Index*, and all other independent variables are defined in Appendix A. All models include year and industry fixed effects. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

	Opt	Bias	Foreca	st_Miss	Асси	ıracy
	(1)	(2)	(3)	(4)	(5)	(6)
Rep_Dum	-0.143†		-0.142†		0.087^{\dagger}	
	(-3.92)		(-3.64)		(4.58)	
Rep_Index		-0.158†		-0.194†		0.084^{\dagger}
		(-3.50)		(-4.09)		(3.60)
Ln(assets)	-0.069†	-0.073†	-0.038**	-0.041**	0.020^{\dagger}	0.022^{\dagger}
	(-4.61)	(-4.84)	(-2.40)	(-2.57)	(2.72)	(3.00)
MB	-0.007**	-0.008**	-0.005	-0.006	0.014^{\dagger}	0.014^{\dagger}
	(-1.96)	(-2.02)	(-1.39)	(-1.44)	(6.66)	(6.71)
Leverage	0.333**	0.331**	0.311**	0.312**	-0.297†	-0.297†
	(2.29)	(2.28)	(2.00)	(2.00)	(-5.24)	(-5.25)
RD	2.555^{\dagger}	2.532^{\dagger}	1.827^{\dagger}	1.797^{\dagger}	-1.433†	-1.424 [†]
	(5.80)	(5.75)	(3.99)	(3.93)	(-8.84)	(-8.78)
ROA	-2.777†	-2.786†	-2.620†	-2.625†	0.675^{\dagger}	0.678^{\dagger}
	(-11.87)	(-11.91)	(-10.87)	(-10.90)	(10.53)	(10.57)
Volatility	17.410^{\dagger}	17.350 [†]	15.227^{\dagger}	15.160^{\dagger}	-15.288†	-15.274†
	(11.20)	(11.17)	(9.38)	(9.34)	(-23.44)	(-23.41)
Ln(Analyst)	-0.458†	-0.459†	-0.409†	-0.409^{\dagger}	0.383^{\dagger}	0.384^{\dagger}
	(-16.63)	(-16.66)	(-13.92)	(-13.92)	(31.87)	(31.90)
Instit_Own	-0.196†	-0.193†	-0.068	-0.065	0.086^{\dagger}	0.084^{\dagger}
	(-4.21)	(-4.13)	(-1.38)	(-1.30)	(3.52)	(3.43)
Litigation	-0.228†	-0.231 [†]	-0.146**	-0.151**	0.062^{*}	0.062^{*}
	(-3.72)	(-3.78)	(-2.30)	(-2.38)	(1.93)	(1.94)
News	-0.849†	-0.849†	-0.886^{\dagger}	-0.886^{\dagger}	0.108^{\dagger}	0.108^{\dagger}
	(-23.27)	(-23.27)	(-22.08)	(-22.09)	(6.69)	(6.70)
Equity_Issue	-0.046	-0.046	0.025	0.024	-0.027	-0.028
	(-1.08)	(-1.08)	(0.54)	(0.53)	(-1.37)	(-1.39)
Acquisition	-0.279†	-0.280†	-0.206†	-0.207†	0.226^{\dagger}	0.226^{\dagger}
	(-8.68)	(-8.70)	(-6.01)	(-6.03)	(13.74)	(13.75)
Industry_Conc	-0.568†	-0.572*	-0.549†	-0.557*	0.458^{\dagger}	0.457^{\dagger}
	(-3.55)	(-3.57)	(-3.24)	(-3.29)	(6.13)	(6.11)
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes
Observations	33,951	33,951	33,951	33,951	33,951	33,951
Pseudo / Adj. R^2	0.220	0.220	0.176	0.176	0.266	0.266

Table 5. CEO Political Ideology: Bad news, Good news, and Earnings Surprise.

This table presents the results of the logistic regression models of the association between CEO political ideology and the credibility of the management earnings forecasts news. The dependent variable in models (1) and (2) is *Bad_News*, which is an indicator variable that equals one if forecast news is negative, and 0 otherwise. The dependent variable in models (3) and (4) is *Good_News*, which is an indicator variable equals one if forecast news is non-negative, and zero otherwise. Where forecast news is the difference between the management earnings forecasts and the most recent mean analyst estimate deflated by the stock price one trading day prior to the management forecast release date. The dependent variable in models (5) and (6) is *Positive_Surprise*, which is an indicator variable that equals one if an earnings surprise is greater than 0.0001, and zero otherwise. The dependent variable in models (7) and (8) is *Negative_Surprise*, which is an indicator variable that equals one if an earnings surprise is less than -0.0001, and zero otherwise. The dependent variable in models (9) and (10) is *Neutral_Surprise*, which is an indicator variable that equals one if an earnings surprise is between 0.0001 and -0.0001, and zero otherwise. Measures of CEO political ideology, *Rep_Dum*, and *Rep_Index*, and all other independent variables are defined in Appendix A. All models include year and industry fixed effects. T-statistics are computed using robust standard errors and reported in parentheses. † , ** and * denote significance at the 1%, 5%, and 10% levels, respectively.

Tespectively.	Bad_News		Good	News	Posi		0	itive_	Neu	
						orise		prise		orise
N D	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
Rep_Dum	0.133†		0.052		0.058**		-0.090†		0.029	
D I I	(3.84)	0 1 1 1	(1.40)	0.011	(2.07)		(-2.97)	0 00 4**	(0.64)	0.06
Rep_Index		0.144 [†]		0.011		0.052		-0.094**		0.065
- ()		(3.35)		(0.23)	**	(1.51)	0.04 0 #	(-2.54)	· · · · · · ·	(1.18)
Ln(assets)	0.052†	0.055†	-0.016	-0.014	-0.023**	-0.022*	0.042†	0.040 [†]	-0.077†	-0.077*
	(3.56)	(3.78)	(-1.07)	(-0.93)	(-2.04)	(-1.92)	(3.32)	(3.19)	(-4.07)	(-4.09)
MB	0.006	0.006^{*}	-0.010**		-0.009†	-0.009†	-0.011†	-0.011*	0.030†	0.030 [†]
	(1.60)	(1.65)	(-2.29)	(-2.26)	(-2.86)	(-2.84)	(-3.20)	(-3.22)	(7.05)	(7.05)
Leverage	0.184	0.184	0.799†	0.801^{\dagger}	-0.409†	-0.410 [†]	0.837^{\dagger}	0.838^{\dagger}	-1.304†	-1.304†
	(1.38)	(1.39)	(5.99)	(6.00)	(-4.07)	(-4.08)	(7.91)	(7.92)	(-6.51)	(-6.51)
RD	-3.800†	-3.781 [†]	-3.119†	-3.120 [†]	1.951†	1.955†	-1.751†	-1.760†	-1.264†	-1.253†
	(-9.16)	(-9.12)	(-7.12)	(-7.12)	(6.79)	(6.81)	(-5.62)	(-5.65)	(-2.67)	(-2.64)
ROA	2.503†	2.511†	0.454**	0.459^{**}	1.392†	1.394†	-1.849†	-1.851†	1.036†	1.036†
	(12.05)	(12.08)	(2.44)	(2.46)	(10.58)		(-13.18)		(4.27)	(4.27)
Volatility				-18.294†		-0.458	2.559**	2.550^{**}	-9.098†	-9.080†
	· · · · ·	· /	· /	(-11.75)	· /	(-0.42)	(2.19)	(2.18)	(-4.88)	(-4.87)
Ln(Analyst)	0.582^{\dagger}	0.583†	0.563†	0.563†	0.353†	0.353†	-0.595†	-0.595†	0.614†	0.614^{\dagger}
	(21.80)	(21.81)	(19.58)	(19.62)	(17.57)		· /	(-27.73)	(16.87)	(16.85)
Instit_Own	0.252^{\dagger}	0.249†	0.323^{\dagger}	0.322^{\dagger}	0.295†	0.294†	-0.313†	-0.311†	0.046	0.044
	(5.74)	(5.66)	(6.78)	(6.77)	(8.25)	(8.22)	(-8.28)	(-8.22)	(0.76)	(0.74)
Litigation	0.277^{\dagger}	0.279^{\dagger}	0.059	0.058	-0.010	-0.010	-0.054	-0.055	0.055	0.057
	(4.73)	(4.77)	(0.92)	(0.90)	(-0.21)	(-0.20)	(-1.07)	(-1.08)	(0.74)	(0.77)
News	0.290^{\dagger}	0.290^{\dagger}	-0.782†	-0.782^{\dagger}	0.505^{\dagger}	0.505^{\dagger}	-0.663†	-0.664†	0.260†	0.260^{\dagger}
	(8.99)	(8.99)		(-23.42)	· · · ·	· /	(-25.04)	· · · · ·	(5.86)	(5.86)
Equity_Issue	0.022	0.022	-0.109**		-0.019	-0.019	0.006	0.006	0.038	0.038
	(0.55)	(0.54)	(-2.49)	(-2.51)	(-0.62)	(-0.62)	(0.17)	(0.17)	(0.77)	(0.78)
Acquisition	0.316†	0.317^{\dagger}	0.279^{\dagger}	0.279^{\dagger}	0.030	0.030	-0.098†	-0.098†	0.195†	0.195†
	(10.40)	(10.42)	(8.57)	(8.58)	(1.24)	(1.24)	(-3.76)	(-3.77)	(4.93)	(4.93)
Industry_Conc	0.835^{\dagger}	0.838^{\dagger}	0.354**	0.350^{**}	-0.370†	-0.371†	0.181	0.182	0.444^{**}	0.447^{**}
	(5.59)	(5.61)	(2.17)	(2.14)	(-3.29)	(-3.30)	(1.52)	(1.53)	(2.43)	(2.45)
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951
Pseudo R ²	0.257	0.257	0.180	0.180	0.052	0.052	0.102	0.102	0.077	0.077

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Table 6. Propensity Score Matching

This table presents the test of management earnings forecasts between Republican and matching samples of control firm-years with non-Republican CEOs matched primarily on the firm characteristics, year, and industry. Panel A presents results for the diagnostic- differences in means of firm characteristics where *Treatment* denotes *Rep_dum_{cycle}* which is an indicator variable that equals one if all donations of a CEO in an election cycle are directed to the Republican Party and *controls* refers to matching sample of CEOs who donated to other parties or never donated. *Difference* represents the difference between treated and control groups. Panel B & C presents the results for the models of the association between management earnings forecasts and CEO political ideology from matched firm-years. All other variables are defined in Appendix A. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

11 /	Panel A: Di	agnostic- diff	erences in m	eans of variable	es	
Variable	Treatment		Control	Dif	ference	T-stat
Ln(assets)	7.620		7.604	0	.015	0.48
MB	3.209		3.250	-(0.041	-0.53
Leverage	0.164		0.160	0	.004	1.55
RD	0.023		0.023	0	.000	-0.31
ROA	0.049		0.051		0.002	-1.09
Return_Volatility	0.025		0.025		0.000	0.40
Ln(Analyst)	2.268		2.264	0	0.004	0.25
Instit_Own	0.558		0.550		.009	1.20
Litigation	0.157		0.151	0	0.005	0.74
News	0.650		0.649		.001	0.15
Equity_Issue	0.156		0.160		0.004	-0.53
Acquisition	0.416		0.412		0.004	0.39
Industry_Conc	0.482		0.480		.002	0.59
Pan	el B. CEO Politi	ical ideology a	and manager	nent earnings fo	precast (1)	
	Issue	Frequency	Range	Ln(Horizon)	OptBias	Forecast_ Miss
	(1)	(2)	(3)	(4)	(5)	(6)
Rep_dum _{cycle}	0.103**	0.091^{*}	0.104^{*}	0.083^{*}	-0.134**	-0.218†
	(2.03)	(1.94)	(1.91)	(1.88)	(-2.39)	(-3.68)
Controls	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes
Observations	9,578	9,578	9,578	9,578	9,578	9,578
Pseudo / Adj. R^2	0.254	0.301	0.269	0.279	0.221	0.176
Pane	el C. CEO Politi	cal ideology a	nd managen	nent earnings fo	precast. (2)	
	Accuracy	Bad_	$Good_{-}$	Positive_	Negative_S	Neutral_
	Accuracy	News	News	Surprise	urprise	Surprise
	(7)	(8)	(9)	(10)	(11)	(12)
Rep_dum _{cycle}	0.071**	0.090^{*}	-0.031	0.071^{*}	-0.096**	0.017
	(2.48)	(1.68)	(-0.55)	(1.68)	(-2.12)	(0.24)
Controls	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes
Observations	9,578	9,578	9,578	9,578	9,578	9,578
Pseudo / Adj. R^2	0.279	0.265	0.168	0.046	0.093	0.093

Table 7. MEF around CEO turnover. Difference-in-Difference (DID) test.

This table presents estimates from the Difference-in-Difference (DID) regressions of the association between CEO political ideology and management earnings forecasts around CEO turnover events (-3, +3). *Rep-Leaving* is an indicator variable equals one if a firm replaces a Rep CEO with a non-Rep CEO, 0 otherwise. Republican CEOs are defined as *Rep_dum_{Only}*, which is an indicator variable that equals one if all donations of a CEO in an election cycle are directed to the Republican Party only. *After* is an indicator variable equals 1 for the years after the CEO turnover, 0 for the pre-tenure period where *CEO turnover* equals one if a CEO in the current year is different from the CEO in the previous year. We only consider turnover events where long-term old CEOs are replaced by long-term new CEOs (long-term old and long-term new CEOs are those who hold their position for at least three years). All models include control variables, year, and industry fixed effects. All other independent variables are defined in Appendix A. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

	Issue	Frequency	Range	Ln	<i>OptBias</i>	Forecast_	Accuracy	Bad_	$Good_{-}$	Positive_	Negative_	Neutral_
	15500	Frequency	Kunge	(Horizon)	OpiDius	Miss	Accuracy	News	News	Surprise	Surprise	Surprise
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Rep-Leaving*After	-0.274**	-0.080	-0.230*	-0.246*	0.410^{\dagger}	0.369**	-0.188**	-0.292**	-0.054	-0.150	0.208^*	-0.134
	(-2.03)	(-0.62)	(-1.67)	(-1.92)	(2.84)	(2.47)	(-2.24)	(-2.12)	(-0.37)	(-1.31)	(1.68)	(-0.73)
Rep-Leaving	0.105	-0.028	-0.026	0.113	-0.177	-0.229*	0.139**	0.121	-0.058	-0.062	0.034	0.077
	(0.96)	(-0.30)	(-0.23)	(1.12)	(-1.51)	(-1.90)	(2.09)	(1.09)	(-0.48)	(-0.67)	(0.34)	(0.53)
After	0.100	0.155^{**}	0.178^{**}	0.114^{*}	-0.275†	-0.243†	0.050	0.187^{**}	-0.103	-0.043	0.072	-0.043
·	(1.44)	(2.43)	(2.34)	(1.72)	(-3.51)	(-2.94)	(1.14)	(2.53)	(-1.36)	(-0.71)	(1.10)	(-0.44)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	8,722	8,722	8,722	8,722	8,722	8,722	8,722	8,722	8,722	8,722	8,722	8,722
<i>Pseudo / Adj.</i> R^2	0.216	0.299	0.214	0.253	0.200	0.147	0.271	0.225	0.150	0.054	0.106	0.081

Table 8. Controlling for CEO characteristics, incentives, and overconfidence.

This table presents tests of the association between CEO political ideology and management earnings forecast controlling for CEO characteristics (Ln(Tenure), Ln(Age), Duality, CEO Gender, Ln(Delta), Ln(Vega), CEO_Own, and Overconfidence) in addition to the baseline control variables. Panel A reports results for Issue, Frequency, Range, Ln(Horizon), OptBias, and Forecast_Miss, in turn. Panel B reports results for Accuracy, Bad_News, Good_News, Positive_Surprise, Negative_Surprise, and Neutral_Surprise, in turn. Measures of CEO political ideology, Rep_Dum and Rep_Index, and all other independent variables are defined in Appendix A. All models include year and industry fixed effects. T-statistics are computed using robust standard errors and reported in parentheses.[†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively

	Panel A. CEO Political ideology and management earnings forecast (1)													
	Iss	sue	Freq	uency	Ra	nge	Ln(He	orizon)	Opt	tBias	M	liss		
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)		
Rep_Dum	0.122^{\dagger}		0.177^{\dagger}		0.116^{\dagger}		0.115^{\dagger}		-0.136†		-0.134†			
	(3.37)		(5.02)		(3.04)		(3.59)		(-3.42)		(-3.16)			
Rep_Index		0.117^{\dagger}		0.126^{\dagger}		0.126^{\dagger}		0.103 [†]		-0.139†		-0.184†		
		(2.62)		(2.91)		(2.69)		(2.61)		(-2.86)		(-3.57)		
Ln(Tenure)	-0.027	-0.025	-0.092^{\dagger}	-0.088^{\dagger}	-0.022	-0.020	-0.052**	-0.050**	0.095^{\dagger}	0.092^{\dagger}	0.089^{\dagger}	0.087^{\dagger}		
	(-1.05)	(-0.96)	(-4.31)	(-4.13)	(-0.80)	(-0.73)	(-2.50)	(-2.39)	(3.30)	(3.22)	(2.91)	(2.87)		
Ln(Age)	-0.130	-0.129	-0.385†	-0.380^{\dagger}	-0.099	-0.100	-0.137	-0.136	0.141	0.142	-0.086	-0.079		
	(-1.00)	(-0.99)	(-3.35)	(-3.30)	(-0.72)	(-0.73)	(-1.24)	(-1.22)	(0.97)	(0.98)	(-0.56)	(-0.51)		
Duality	0.213^{\dagger}	0.215^{\dagger}	0.181^{\dagger}	0.186^{\dagger}	0.211^{+}	0.213^{\dagger}	0.170^{\dagger}	0.172^{\dagger}	-0.156†	-0.158†	-0.140^{\dagger}	-0.141 [†]		
	(6.32)	(6.39)	(5.68)	(5.84)	(5.99)	(6.04)	(5.76)	(5.84)	(-4.20)	(-4.27)	(-3.55)	(-3.56)		
CEO Gender	0.193**	0.190^{**}	0.170	0.163	-0.107	-0.110	0.184^{*}	0.181^{*}	-0.078	-0.074	-0.061	-0.058		
	(2.05)	(2.02)	(1.63)	(1.56)	(-1.08)	(-1.11)	(1.96)	(1.92)	(-0.76)	(-0.72)	(-0.56)	(-0.54)		
Ln(Delta)	0.036^{*}	0.037^{*}	0.073^{\dagger}	0.074^{\dagger}	0.004	0.004	0.043**	0.044**	-0.124 [†]	-0.124†	-0.112^{\dagger}	-0.111 [†]		
	(1.69)	(1.70)	(3.95)	(4.03)	(0.19)	(0.19)	(2.40)	(2.43)	(-5.11)	(-5.13)	(-4.34)	(-4.32)		
Ln(Vega)	0.084^{\dagger}	0.083^{\dagger}	0.096^{\dagger}	0.095^{\dagger}	0.077^{\dagger}	0.077^{\dagger}	0.092^{\dagger}	0.092^{\dagger}	-0.049^{\dagger}	-0.049^{\dagger}	-0.044^{\dagger}	-0.044^{\dagger}		
	(6.59)	(6.57)	(7.49)	(7.46)	(5.87)	(5.85)	(7.93)	(7.91)	(-3.60)	(-3.57)	(-3.03)	(-3.03)		
CEO_Own	-2.692†	-2.666†	-2.210^{\dagger}	-2.197†	-2.790^{\dagger}	-2.759†	-2.120 [†]	-2.105^{\dagger}	3.460^{\dagger}	3.424^{\dagger}	3.559 [†]	3.511 [†]		
	(-5.68)	(-5.63)	(-6.05)	(-6.01)	(-5.28)	(-5.22)	(-5.69)	(-5.65)	(6.07)	(6.01)	(5.69)	(5.62)		
Overconfidence	0.101^{+}	0.102^{+}	0.175^{\dagger}	0.177^{\dagger}	0.085^{**}	0.087^{**}	0.116^{\dagger}	0.117^{\dagger}	-0.062	-0.063	-0.056	-0.058		
	(2.75)	(2.78)	(5.39)	(5.43)	(2.20)	(2.23)	(3.69)	(3.72)	(-1.50)	(-1.53)	(-1.27)	(-1.31)		
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes		
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes		
Observations	27,912	27,912	27,912	27,912	27,912	27,912	27,912	27,912	27,912	27,912	27,912	27,912		
Pseudo R^2/Adj . R^2	0.259	0.259	0.292	0.291	0.254	0.254	0.276	0.276	0.224	0.224	0.178	0.178		

	Acci	uracy	Bad_	News	Good	News	Positive_	Surprise	Negative	Surprise	Neutral	Surprise
	(13)	(14)	(15)	(16)	(17)	(18)	(19)	(20)	(21)	(22)	(23)	(24)
Rep_Dum	0.077^{\dagger}		0.112^{\dagger}		0.056		0.045		-0.064*		0.010	
	(3.67)		(2.97)		(1.41)		(1.43)		(-1.91)		(0.20)	
Rep_Index		0.070^{\dagger}		0.112**		0.016		0.010		-0.047		0.068
		(2.72)		(2.39)		(0.33)		(0.26)		(-1.13)		(1.13)
Ln(Tenure)	-0.040^{\dagger}	-0.038†	-0.050^{*}	-0.048^{*}	0.034	0.036	-0.088^{\dagger}	-0.087^{\dagger}	0.087^{\dagger}	0.086^{\dagger}	0.019	0.018
	(-3.01)	(-2.89)	(-1.86)	(-1.78)	(1.20)	(1.25)	(-4.28)	(-4.22)	(3.98)	(3.91)	(0.55)	(0.54)
Ln(Age)	-0.125*	-0.124*	-0.092	-0.092	-0.191	-0.187	-0.256**	-0.252**	0.346^{\dagger}	0.343 [†]	-0.107	-0.114
	(-1.80)	(-1.79)	(-0.66)	(-0.67)	(-1.33)	(-1.30)	(-2.38)	(-2.35)	(3.00)	(2.99)	(-0.62)	(-0.66)
Duality	0.098^{\dagger}	0.099^{\dagger}	0.205^{\dagger}	0.207^{\dagger}	0.161^{+}	0.164^{\dagger}	0.057^{**}	0.059**	-0.043	-0.044	-0.050	-0.052
	(5.23)	(5.30)	(5.81)	(5.87)	(4.39)	(4.47)	(1.98)	(2.06)	(-1.39)	(-1.44)	(-1.07)	(-1.12)
CEO Gender	-0.017	-0.020	0.095	0.092	0.124	0.121	-0.140^{*}	-0.142*	0.125	0.128	0.090	0.092
	(-0.30)	(-0.35)	(0.98)	(0.94)	(1.27)	(1.24)	(-1.65)	(-1.68)	(1.39)	(1.42)	(0.65)	(0.67)
Ln(Delta)	0.066^{\dagger}	0.066^{\dagger}	0.084^{\dagger}	0.085^{\dagger}	-0.045*	-0.045*	0.097^{\dagger}	0.098^{\dagger}	-0.177^{\dagger}	-0.177^{\dagger}	0.159†	0.159†
	(5.70)	(5.74)	(3.75)	(3.76)	(-1.95)	(-1.91)	(5.48)	(5.52)	(-9.25)	(-9.28)	(5.38)	(5.35)
Ln(Vega)	0.053^{\dagger}	0.053^{\dagger}	0.065^{\dagger}	0.065^{\dagger}	0.071^{\dagger}	0.071^{\dagger}	-0.001	-0.002	0.015	0.015	-0.023	-0.023
	(7.04)	(7.03)	(5.00)	(4.98)	(4.86)	(4.84)	(-0.13)	(-0.15)	(1.25)	(1.27)	(-1.34)	(-1.31)
CEO_Own	-1.845†	-1.835†	-3.248†	-3.220†	-0.717	-0.713	-1.211 [†]	-1.212†	2.302^{\dagger}	2.297^{\dagger}	-2.197†	-2.182†
	(-7.90)	(-7.85)	(-6.13)	(-6.08)	(-1.40)	(-1.39)	(-3.28)	(-3.28)	(5.93)	(5.91)	(-3.40)	(-3.38)
Overconfidence	0.088^{\dagger}	0.089^{\dagger}	0.107^{\dagger}	0.108^{\dagger}	0.077^{*}	0.077^{*}	0.044	0.045	-0.050	-0.050	-0.008	-0.008
	(4.47)	(4.50)	(2.74)	(2.77)	(1.87)	(1.87)	(1.45)	(1.46)	(-1.54)	(-1.55)	(-0.16)	(-0.15)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	27,912	27,912	27,912	27,912	27,912	27,912	27,912	27,912	27,912	27,912	27,912	27,912
Pseudo R^2/Adj . R^2	0.278	0.278	0.259	0.259	0.176	0.176	0.0539	0.0539	0.103	0.103	0.0746	0.0746

Panel B. CEO Political ideology and management earnings forecast (2)

Table 9. Alternative measures of CEO Republican Ideology.

This table presents tests of the association between CEO political ideology and management earnings forecast using alternative measures of CEO political ideology. Rep_index_{cycle} is an index calculated as total donations to the Republican Party minus total donations to the Democratic Party divided by total donations to both parties in each election cycle. Rep_dum_{cycle} is an indicator variable that equals one if all donations of a CEO in an election cycle are directed to the Republican Party. Rep_dum_{tenure} is an indicator variable that equals one if all donations of a CEO during her/his tenure are directed to the Republican Party. Panel A reports results for *Issue, Frequency, Range, Ln(Horizon), OptBias, and Forecast_Miss,* in turn. Panel B reports results for *Accuracy, Bad_News, Good_News, Positive_Surprise, Negative_Surprise,* and *Neutral_Surprise,* in turn. All models include control variables, year, and industry fixed effects. All other independent variables are defined in Appendix A. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

0	,			Panel	A. Ált	ernative	e measi	ures of	CEO Re	epublic	an Idec	ology (1)					
		Issue		F_{i}	requen	су		Range		Ln	ı(Horiz	on)		OptBic	as	For	recast_l	Miss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)	(17)	(18)
<i>Rep_index</i> _{cycle}	0.078^{\dagger}			0.088^{\dagger}			0.104†			0.069†			- 0.081*	*		-0.100	t	
	(2.73)			(3.16)			(3.42)			(2.71)			(-2.56)		(-2.94)	
Rep_dum _{cycle}		0.112^{\dagger}			0.099†			0.101**			0.087^{**}	¢		-0.120)†		-0.161	ŕ
		(2.88)			(2.69)			(2.47)			(2.55)			(-2.80)		(-3.55))
Rep_dum _{tenure}			0.122†			0.110†			0.108**			0.087**	k		- 0.097*	*		-0.141†
			(2.76)			(2.65)			(2.33)			(2.29)			(-2.00))		(-2.75)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	33,951	33,951	33,951	· ·	,	· ·	,	· ·	,	·	· ·	,	,	,	133,951	· ·	,	,
<i>Pseudo / Adj.</i> R^2	0.257	0.257	0.257											0.220	0.220	0.176	0.176	0.176
								ures of (<u> </u>		.	,					
	1	4ccurac	cy	В	ad_Net	WS	G	ood_Ne	ews	Posit	tive_Su	rprise	Nega	tive_Sı	ırprise	Neut	ral_Sur	prise
<i>Rep_index</i> _{cycle}	0.046^{\dagger}			0.060**	k		0.007			0.048**	*		- 0.048*	k		-0.015		
	(2.76)			(1.98)			(0.20)			(1.96)			(-1.81))		(-0.38)		
Rep_dum _{cycle}		0.069†			0.111^{\dagger}			0.009			0.039			-0.055			0.019	
		(3.08)			(2.73)			(0.20)			(1.18)			(-1.54)			(0.35)	
Rep_dum _{tenure}			0.070^{\dagger}			0.128†			0.057			0.039			-0.068*			0.058
			(2.82)			(2.77)			(1.15)			(1.05)			(-1.70)			(0.97)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	,	,	· ·	· ·	,	,	,	,	,	,	,	,	,	,	33,951	,	,	· ·
<i>Pseudo / Adj.</i> R^2	0.266	0.266	0.266	0.257	0.257	0.257	0.180	0.180	0.180	0.052	0.052	0.052	0.102	0.102	0.102	0.077	0.077	0.077

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Table 10. Cross-sectional test: High vs. low institutional ownership.

This table presents results for firms with high (above-median) level of institutional ownership (Panel A) and firms with low (below-median) level of institutional ownership (Panel B). Panel A reports results for *Issue, Frequency, Range, Ln(Horizon), OptBias, and Forecast_Miss, in turn.* Panel B reports results for *Accuracy, Bad_News, Good_News, Positive_Surprise, Negative_Surprise, and Neutral_Surprise, in turn.* Measures of CEO political ideology, *Rep_Dum* and *Rep_Index,* and all other independent variables are defined in Appendix A. All models include control variables, year, and industry fixed effects. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

-				Panel A1.	High instit	utional ow	nership (1)					
	Iss	ue	Freq	uency	Ra	nge	Ln(Ho	rizon)	Ор	otBias	Forece	ast_Miss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Rep_Dum	0.177^{\dagger}		0.245^{\dagger}		0.202^{\dagger}		0.162^{\dagger}		-0.185†		-0.170†	
	(3.94)		(4.97)		(4.40)		(3.83)		(-3.84)		(-3.32)	
Rep_Index		0.138**		0.192†		0.182^{\dagger}		0.134†		- 0.166 [†]		-0.213†
		(2.54)		(3.21)		(3.28)		(2.62)		(-2.84)		(-3.47)
Observations	16,975	16,975	16,975	16,975	16,975	16,975	16,975	16,975	16,971	16,971	16,971	16,971
Pseudo /Adj. R ²	0.221	0.220	0.269	0.268	0.207	0.206	0.251	0.251	0.193	0.193	0.151	0.151
				Panel A2.	High instit	utional ow	nership (2)					
	Асси	ıracy	Bad	News	Good	News	Positive	Surprise	Negative	Surprise	Neutral	Surprise
Rep_Dum	0.141^{\dagger}		0.194†		0.091*		0.077^{*}		-0.058		-0.092	
	(4.98)		(4.19)		(1.87)		(1.93)		(-1.35)		(-1.42)	
Rep_Index		0.138†		0.163 [†]		0.016		0.044		-0.030		-0.072
		(4.02)		(2.89)		(0.27)		(0.92)		(-0.59)		(-0.90)
Observations	16,975	16,975	16,975	16,975	16,975	16,975	16,975	16,975	16,975	16,975	16,971	16,971
Pseudo /Adj. R ²	0.263	0.262	0.223	0.223	0.140	0.139	0.036	0.036	0.074	0.074	0.075	0.075
				Panel B1.	Low instit	utional ow	nership (1)					
	Iss	ue	Freq	uency	Ra	nge	Ln(Ho	rizon)	Opt	Bias	Foreca	st_Miss
Rep_Dum	0.091^{*}		0.073^{*}		0.019		0.066^{*}		-0.098^{*}		-0.143**	
	(1.78)		(1.95)		(0.33)		(1.71)		(-1.68)		(-2.31)	
Rep_Index		0.147^{**}		0.022		0.106		0.073		-0.157**		-0.201 [†]
		(2.25)		(0.49)		(1.48)		(1.50)		(-2.13)		(-2.58)
Observations	16,970	16,970	16,976	16,976	16,970	16,970	16,976	16,976	16,915	16,915	16,915	16,915
Pseudo /Adj. R ²	0.291	0.291	0.293	0.293	0.302	0.303	0.279	0.279	0.245	0.245	0.208	0.208

				Panel B2.	Low institu	utional ow	nership (2)					
	Асси	ıracy	Bad_	News	Good	News	Positive_	Surprise	Negative	Surprise	Neutral_	Surprise
Rep_Dum	0.037		0.076		-0.000		0.052		-0.126†		0.128**	
	(1.49)		(1.39)		(-0.00)		(1.27)		(-2.90)		(2.02)	
Rep_Index		0.028		0.167^{**}		0.032		0.067		-0.166†		0.205^{\dagger}
		(0.91)		(2.40)		(0.42)		(1.32)		(-3.05)		(2.62)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	16,976	16,976	16,927	16,927	16,970	16,970	16,976	16,976	16,976	16,976	16,970	16,970
Pseudo R^2/Adj . R^2	0.266	0.266	0.281	0.281	0.226	0.226	0.067	0.067	0.125	0.126	0.09	0.091

Table 10. Cross-sectional test: High vs. low institutional ownership. Cont'd

Table 11. Cross-sectional test: High vs. low litigation risk.

This table presents results for firms in industries with high litigation environment (Panel A) and firms in industries with low litigation environment (Panel B). Panel A reports results for *Issue, Frequency, Range, Ln(Horizon), OptBias, and Forecast_Miss, in turn.* Panel B reports results for *Accuracy, Bad_News, Good_News, Positive_Surprise, Negative_Surprise,* and *Neutral_Surprise,* in turn. Measures of CEO political ideology, *Rep_Dum* and *Rep_Index,* and all other independent variables are defined in Appendix A. All models include control variables, year, and industry fixed effects. T-statistics are computed using robust standard errors and reported in parentheses.[†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

				Panel A	1. High lit	igation env	vironment (1)				
	Iss	sue	Freq	uency	Ra	nge	Ln(Ha	orizon)	Ор	tBias	Forec	ast_Miss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Rep_Dum	0.220^{\dagger}		0.325^{\dagger}		0.343^{\dagger}		0.215^{\dagger}		-0.227^{\dagger}		-0.063	
	(2.79)		(4.03)		(4.14)		(3.11)		(-2.67)		(-0.67)	
Rep_Index		0.261 [†]		0.318^{\dagger}		0.354^{\dagger}		0.225^{\dagger}		-0.184*		-0.129
		(2.76)		(3.37)		(3.53)		(2.69)		(-1.74)		(-1.11)
Observations	8,120	8,120	8,144	8,144	8,120	8,120	8,144	8,144	8,120	8,120	8,120	8,120
Pseudo /Adj. R ²	0.253	0.253	0.288	0.287	0.223	0.222	0.269	0.269	0.219	0.219	0.181	0.181
				Panel A	2. High lit	igation env	vironment (2)				
	Асси	ıracy	Bad	News	Good	News	Positive	Surprise	Negative	Surprise	Neutral	Surprise
Rep_Dum	0.206^{\dagger}		0.194**		0.149*		0.035		-0.075		0.024	
*	(4.35)		(2.37)		(1.68)		(0.50)		(-0.94)		(0.26)	
Rep Index		0.229†	. ,	0.170^{*}	. ,	0.183*	. ,	0.095		-0.129	. ,	0.030
		(4.00)		(1.71)		(1.68)		(1.11)		(-1.35)		(0.24)
Observations	8,144	8,144	8,120	8,120	8,120	8,120	8,144	8,144	8,144	8,144	8,144	8,144
Pseudo /Adj. R^2	0.272	0.272	0.252	0.252	0.143	0.143	0.048	0.049	0.100	0.100	0.061	0.061
				Panel B	1. Low liti	gation env	ironment (1)				
	Iss	sue	Freq	uency	Ra	nge	Ln(Ho	orizon)	Opti	Bias	Foreca	st Miss
Rep_Dum	0.097^{\dagger}		0.112†		0.071^{*}		0.075**		-0.116 [†]		-0.151†	
1	(2.63)		(3.33)		(1.83)		(2.38)		(-2.87)		(-3.53)	
Rep Index	```	0.075^{*}	` /	0.042	. ,	0.080^{*}	. /	0.048	. /	-0.130 [†]	` '	-0.189†
· _		(1.65)		(1.01)		(1.65)		(1.23)		(-2.61)		(-3.61)
Observations	25,807	25,807	25,807	25,807	25,807	25,807	25,807	25,807	25,807	25,807	25,807	25,807
Pseudo /Adj. R^2	0.266	0.266	0.290	0.290	0.268	0.268	0.281	0.280	0.223	0.223	0.178	0.178

				Panel B	2. Low liti	gation env	ironment (1	2)				
	Accuracy		Bad News		Good News Pos		Positive	Positive Surprise		Negative Surprise		Surprise
Rep_Dum	0.054^{\dagger}		0.107†		0.022		0.068^{**}		-0.097†		0.021	
	(2.62)		(2.79)		(0.54)		(2.19)		(-2.95)		(0.41)	
Rep Index	. ,	0.038		0.114^{**}	. ,	-0.037	. ,	0.046	. ,	-0.092**		0.066
		(1.50)		(2.38)		(-0.72)		(1.19)		(-2.26)		(1.05)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	25,807	25,807	25,807	25,807	25,807	25,807	25,807	25,807	25,807	25,807	25,807	25,807
Pseudo /Adj. R^2	0.275	0.275	0.262	0.262	0.190	0.190	0.053	0.053	0.099	0.099	0.083	0.083

Table 11. Cross-sectional test: High vs. low litigation risk. Cont'd

Internet Appendix

Introduction

This online supplementary material complements and extends our main analysis in "Political orientation and the information revelation preferences of Red and Blue CEOs" in multiple ways. First, we use alternative measures of CEO political ideology and management earnings forecasts to mitigate measurement error. Second, we conduct a range of robustness tests to address various specification issues that could otherwise confound our main results. Lastly, we conduct additional tests to address further the endogeneity issues that could arise from measurement error, selection bias, and/or omitted variable bias.

1. Alternative measures of CEO political ideology and overconfidence

Table A1 presents the results using two alternative measures of Republican ideology, *Rep_indexyear*, which is an index calculated as total donations to the Republican Party minus total donations to the Democratic Party divided by total donations to both parties in each fiscal year, and *Rep_indextenure*, which is an index calculated as total donations to the Republican Party minus total donations to the Democratic Party divided by total donations to the Republican Party minus total donations to the Democratic Party divided by total donations to both parties in a CEO's entire tenure. In Table A2, we use alternative measures of Democratic ideology and Other ideology. Results are similar to our main findings, which mitigate the concerns that our findings are sensitive to our baseline measures of Republican ideology. Further, Table A3 uses *Net_buyer* as an alternative measure of CEO overconfidence in addition to other CEO and firm characteristics.

2. Controlling for variations in CEO donation

Political ideology data includes a significant variation in CEO donation. While some CEOs consistently donate in each election cycle, others never make any political donations. To make

1

sure that such variation does not affect our baseline results, we run a subsample analysis after excluding CEOs who never donated during the sample period (Table A4, Panels A & B). Further, we run a subsample analysis by restricting the sample to CEO donation years only (Table A4, Panels C & D). Results using these restrictive subsamples are similar to our baseline results.

3. Propensity score matching. Using alternative measures of CEO political ideology.²⁶

We run our PSM using alternative measures of CEO political ideology. First, we identify *Treatment* using *Rep_dumonly*, which is an indicator variable that equals one if all donations of a CEO in an election cycle are directed to the Republican Party only (neither Democratic nor others) and *control* refer to a matching sample of CEOs who donated to other parties or never donated (Table A5, Panel A). Next, we identify *Treatment* using *Rep_dum_{cycle}*, which is an indicator variable that equals 1 if all donations of a CEO in an election cycle are directed to the Republican Party, and *control* refers to matching sample if the donations of a CEO in an election cycle are all directed toward the Democratic Party (Table A5, Panel B). Lastly, we identify *Treatment* using *Rep_dum_{tenure}*, which is an indicator variable that equals 1 if all donations variable that equals 1 if all donations of a CEO in an election cycle are all directed toward the Democratic Party (Table A5, Panel B). Lastly, we identify *Treatment* using *Rep_dum_{tenure}*, which is an indicator variable that equals 1 if all donations of a CEO during her/his entire tenure are directed to the Republican Party and *control* refers to matching sample if all donations of a CEO during her/his entire tenure are directed to the Republican Party and *control* refers to matching sample if all donations of a CEO during her/his entire tenure are directed to the Democratic Party (Table A5, Panel C). We carefully match the *Treatment* and *Control* groups on multiple firm characteristics as well as year and industry to mitigate the endogeneity issue. The results of these tests are similar to our baseline PSM results.

²⁶ We do not report the diagnostic tests for the difference in mean matching variables between treatment and control groups for brevity. These tests are available upon request.

4. Management earnings forecasts around CEO turnover.

Our baseline DID test uses a -3, +3 window around CEO turnover events. To address the possibility that our DID results are affected by the window selection, we repeat our DID test using a -2, +2 window, and report results in Table A6. *After* is an indicator variable equals 1 for the years after the CEO turnover. We only consider turnover events where long-term old CEOs are replaced by long-term new CEOs (long-term is defined as holding the position for at least two years). *Rep_Leaving* is an indicator variable equals one if a firm replaces a Rep CEO with a non-Rep CEO, 0 otherwise. Republican CEOs are defined using *Rep_dum_{Only}*, which is an indicator variable that equals one if all donations of a CEO in an election cycle are directed to the Republican Party only (neither Democratic nor others). Consistent with our baseline results, the coefficient of *After*Rep_Leaving is* significantly negative in models of MFE *Issue, Frequency, Range, Horizon, Accuracy, Bad_News, and Positive_Surprise* and significantly positive in models of *Forecast Miss, Bias,* and *Negative Surprise*.

Next, we examine the effect of change in CEO political ideology due to CEO turnover on change on earnings forecasts (Table A7). Specifically, $\Delta dependent$ is the difference between the first full fiscal year under the new CEO and the last full fiscal year under the old CEO. ΔREP_{CEO} is defined as the changes in CEO political ideology due to CEO turnover, where $\Delta REP_{CEO} = 1$ if a Republican CEO replaces a Democratic CEO, 0 if CEO political ideology does not change with turnover, and -1 if a Democratic CEO replaces a Republican CEO.²⁷ Even though this test uses a significantly smaller sample size, the results are similar to our baseline results.

²⁷ Due to this restrictive definition of changes in CEO political ideology measures around CEO turnover event, our sample size is reduced significantly.

5. Additional robustness checks.

To rule out the possibility that our results are not persistent beyond CEO turnover years, we exclude firm-years in which CEO turnover occurred (Table A8, Panel A). To further check the persistency of our baseline results, we exclude the first three years of CEO tenure (Table A8, Panel B). Following Chava, Livdan, and Purnanandam (2009), we use change-on-change regressions to examine the active managerial influence on management earnings forecasts. Specifically, we estimate annual changes in all management earnings forecasts variables, key Republican measures, and control variables similar to our baseline regressions. Following Hutton, Jiang, and Kumar (2014), we restrict our sample to those firm-years where annual changes in both Republican measures and management earnings forecast variables are non-zero. The results of this test are presented in Table A9. Further, Baik, Farber, and Lee (2011) find that CEO ability is positively associated with the likelihood, frequency, and accuracy of earnings forecasts. Republican CEOs may have higher ability compared to non-Republican CEOs driving our main findings. Thus, following Demerjian, Lev, and McVay (2012), we control for managerial ability and find similar results to our baseline regression.²⁸ The results of this test are presented in Table A10.

To capture the state-level variations in CEO political ideology and management earnings forecasts, we control for state fixed effects (headquarters) (table A11, Panel A & B). Moreover, our results continue to hold if we cluster the standard error at the firm level (Table A11, Panel C & D).

²⁸ Thank you Demerjian, Lev, and McVay (2012) for sharing their data. Managerial ability data is available at: <u>https://faculty.washington.edu/pdemerj/data.html</u>. Last accessed on May 24, 2020.

Our baseline results suggest a positive association between CEOs' conservative political ideology (Republican) and the quality of earnings forecast. However, political activism can represent an alternative explanation of our ideology interpretation of the results. To address this issue, we estimate models that concurrently control for CEO's Republican as well as Democratic ideologies. (Table A12). Coefficient estimates of measures of Republican and Democratic ideologies are opposite, which is consistent with the ideology rather than the activism explanation of our results.

Moreover, we restrict our samples to firms that appear at least once in the I/B/E/S to address the database coverage issue (Chuk, Matsumoto, and Miller (2013); Houston et al. (2019)). Specifically, we exclude those firms that have never issued any earnings forecast during our sample period. This setting should eliminate the possible bias in our results caused by the effect of firms that have never issued any EPS forecasting in our sample period. The results of this test are presented in Table A13.

Table A14 presents our results for subsamples of firms with a high and low level of institutional ownership. Table A15 presents our results for a subsample of pre-crisis observations (1993-2007) and a subsample of post-crisis observations (2010-2016).

Lastly, to further ensure a positive relationship observed between CEO political ideology and MEF, we now examine the association between CEO Democratic ideology and the likelihood and properties of MEF. Specifically, we run our models using the Dem_Dum and Dem_Index. The results of these tests are presented in Table A16.

[Insert Table A16]

Results in Table A16 lend strong support to the main premises of this paper. CEO Democratic ideology is negatively associated with forecast issue, frequency, range, horizon, accuracy, bad news, and positive earnings surprise and positively associated with forecast bias, forecast miss, and negative surprise, albeit some of these effects are not statistically significant. Specifically, these results show, on average, that Democratic CEOs are around 8.8% less likely to issue forecasts, compared to CEOs with other political ideologies (model 1). Further, on average, Democratic CEOs are 9% to 12% less likely to miss forecast, 3.6% to 7.4% more likely to experience negative earnings surprise, 4.5% to 7.7% less likely to experience positive earnings surprise, and have 2.5% to 4.3% higher forecast accuracy, compared to non-Democratic CEOs.

Lastly, we attempt to refute the possibility that our results are driven by the large number of non-forecast years in our sample. We run our models using a subsample that excludes firm/year observations with no management earnings forecast. The results of this test are reported in Table A17 and are similar to our baseline results.

{Insert Table A17]

References

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Variable	Definition
CEO political ideo	logy (Baseline)
Rep_Dum	An indicator variable that equals one if a CEO donated more to the Republican party than to the Democratic party during her/his entire tenure [Bhandari et al. 2018].
Rep_Index	The percentage of a CEO's support for the Republican Party calculated as the number of cycles in which a CEO donates exclusively to the Republican Party divided by her/his number of donation cycles in the sample period [Hong and Kostovetsky, 2012].
CEO political ideo	
Dem_Dum	An indicator variable that equals one if a CEO donated more to the Democratic Party than to the Republican Party during her/his entire tenure.
Dem_Index	The percentage of a CEO's support for the Democratic Party calculated as the number of cycles in which a CEO donates exclusively to the Democratic Party divided by her/his number of donation cycles in the sample period.
<i>Rep_dum</i> _{cycle}	An indicator variable that equals one if all donations of a CEO in an election cycle are directed to the Republican Party [Hutton et al. 2014].
Rep_dum _{tenure}	An indicator variable that equals one if all donations of a CEO during her/his entire tenure are directed to the Republican Party [Elnahas and Kim, 2017].
<i>Rep_index</i> _{cycle}	An index calculated as total donations to the Republican Party minus total donations to the Democratic Party divided by total donations to both parties in each election cycle. This index ranges between -1 (strong Democrat) and 1 (strong Republican) [Hutton et al. 2014].
Rep_dum _{Only}	An indicator variable that equals one if all donations of a CEO in an election cycle are directed to the Republican Party only (neither Democratic nor others).
CEO political ideo	logy (Internet appendix)
Rep_index _{year}	An index calculated as total donations to the Republican Party minus total donations to the Democratic Party divided by total donations to both parties in each fiscal year. This index ranges between -1 (strong Democrat) and 1 (strong Republican).
<i>Rep_index</i> _{temure}	An index calculated as total donations to the Republican Party minus total donations to the Democratic Party divided by total donations to both parties in a CEO's entire tenure. This index ranges between -1 (strong Democrat) and 1 (strong Republican).
Dem_dum _{cycle}	An indicator variable that equals one if the donations of a CEO in an election cycle are all directed toward the Democratic Party.
Dem_dum _{tenure}	An indicator variable that equals one if all donations of a CEO during her/his entire tenure are directed to the Democratic Party.
Dem_dum _{cycle2}	An indicator variable that equals one if the donations of a CEO in an election cycle are all directed toward the Democratic Party but not the Republican Party.
Other_Index	Percentage of a CEO's support for other Parties calculated as the number of cycles in which a CEO donates exclusively to other parties divided by her/his total number of donation cycles in the sample period.
Other_dum _{cycle}	An indicator variable that equals one if the donations of a CEO in an election cycle are all directed toward the other parties (neither Republican nor Democratic).

Appendix A. Variable definition

Appendix A. Variable definition- Cont'd

Voluntary disclosure	
Issue	An indicator variable that equals one if a firm makes annual earnings forecasts in
F	a fiscal year.
Frequency	The total number of annual earnings forecasts made by a firm in a fiscal year.
Ln(Horizon)	The natural logarithm of one plus the average horizon of annual earnings forecasts made by a firm in a fiscal year. For each forecast, the horizon is defined as the
	number of calendar days between the forecast announcement date and the corresponding period end date. We assign a value of zero when a firm makes no forecasts in a fiscal year.
Range	An indicator variable of range estimates. For each forecast, we first assign 1 for
lunge	range estimates and zero otherwise. This indicator variable is then averaged for each firm-year. The Range is then defined as an indicator variable that equals one
	if the average range is greater than 0.5, and zero otherwise.
Forecast_Miss	An indicator variable that equals one if a firm misses at least one forecast in a year, and zero otherwise. Where <i>Miss</i> equals one if the actual earning is less than the
	management earnings forecasts, and zero otherwise for open-ended and point forecasts, and equals one if the actual earnings are less than the lower bound of the
	range forecast for range estimates.
OptBias	An indicator variable that equals one if the average <i>Bias</i> in a year is positive, and zero otherwise. Where, for each estimate, <i>Bias</i> is the difference between
	management earnings forecasts and actual earnings scaled by the stock price at the end of the month prior to the forecast.
Accuracy	The average Forecast accuracy for all annual earnings forecasts made by a firm in
ieeu uey	a fiscal year. For each estimate, we first calculate the absolute difference between
	management earnings forecasts and actual earnings scaled by the stock price at the
	end of the month prior to the forecast. Next, we identify forecast accuracy as the
	quintile ranking of the scaled difference, where one is assigned for the top quintile
	(largest error), and five is assigned for the bottom quintile (lowest error), and zero if no forecasts are made.
Bad_News	An indicator variable that equals one if forecast news is negative, and zero
	otherwise. Where forecast news is the difference between the management earnings forecasts and the most recent mean analyst estimate deflated by the stock price one trading day prior to the management forecast release date.
Good News	An indicator variable equals one if forecast news is non-negative, and zero
00000_1100003	otherwise. Forecast news is the difference between the management earnings
	forecasts and the most recent mean analyst estimate scaled by the stock price one
	trading day prior to the management forecast release date.
Positive Surprise	An indicator variable that equals one if an earnings surprise is greater than 0.0001,
osuive_suiprise	and zero otherwise. Earnings surprise is calculated as the difference between the
	actual earnings and the mean analyst estimate scaled by the stock price three
	trading days prior to an earnings announcement.
Negative Surprise	An indicator variable that equals one if an earnings surprise is less than -0.0001,
	and zero otherwise. Earnings surprise is calculated as the difference between the
	actual earnings and the mean analyst estimate scaled by the stock price three
	actual calmings and the mean analyst estimate scaled by the stock blice the

Appendix A. Varial Neutral Surprise	ble definition- Cont'd An indicator variable that equals one if an earnings surprise is between 0.0001 and
weun un_surprise	-0.0001, and zero otherwise. Earnings surprise is calculated as the difference
	between the actual earnings and the mean analyst estimate scaled by the stock price
	three trading days prior to an earnings announcement.
Firm Characteristics	three trading days prior to an earnings announcement.
<i>Ln(assets)</i>	The natural logarithm of total assets (at).
MB	The ratio of market-to-book value of equity. [(prcc_f*csho) / ceq].
Leverage	The ratio of total debt divided by market value of total assets. [(Dltt+Dlc) / (at-
Leverage	ceq+csho*prcc f)].
RD	Expenditures on research and development scaled by total assets. [xrd/at]
ROA	Return on assets measured as income before extraordinary items scaled by total
	assets. [ib/at]
Volatility	The standard deviation of daily stock return (CRSP variable ret) of a firm over the
,	last fiscal year.
Ln(Analyst)	The natural logarithm of the number of analysts following a firm.
Institutional Own	The percentage of shares owned by institutional investors.
Litigation	An indicator variable that equals one if a firm's SIC code is in industries subject
0	to increased litigation (2833-2836, 3570-3577, 3600-3674, and 7370-7374), and
	zero otherwise.
News	An indicator variable that equals one if the current period EPS is greater than or
	equal to the previous-period EPS, and zero otherwise.
Equity_Issue	An indicator variable that equals one if a firm issued shares in a year.
Acquisition	An indicator variable that equals one if a firm's annual acquisitions or merger-
	related costs exceeded five percent of net income (loss) in year t, and zero
	otherwise. [aqc/ni]
Industry_Conc	A firm's industry concentration, measured as the sum of sales of the top five firms
	in its two-digit SIC code scaled by total sales of all firms in its two-digit SIC code
	in year t. $\left[\sum_{i=1}^{5} Sale_{i,j} / \sum_{i=1}^{n} Sale_{i,j}\right]$
CEO Characteristic	CS
Ln(Tenure)	The natural logarithm of CEO tenure, where tenure is defined as the length of a
	CEO's tenure with her/his current firm (measured as fiscal year minus year joined
	as CEO).
Ln(Age)	The natural logarithm of the age of a CEO as of the year in which a management
	earnings forecast was released.
Duality	An indicator variable that equals one if a CEO is also the chairman, and zero
	otherwise.
Ln(Delta)	The natural logarithm of the expected dollar changes in CEO wealth for a 1%
	change in stock price computed as in Core and Guay (2002).
Ln(Vega)	The natural logarithm of the expected dollar changes in CEO wealth for a 1%
	change in stock return volatility computed as in Guay (1999).
CEO_Own	The percentage of shares outstanding owned by a CEO. [SHROWN_EXCL_OPTS
~~~ <i>~ `</i>	/(CSHO*1000)]
CEO Gender	CEO Gender equals 1 if a CEO is female, 0 otherwise.

# Appendix A. Variable definition- Cont'd

# Appendix A. Variable definition- Cont'd

Holder67	An indicator variable that equals one if a CEO holds vested options with average
	moneyness greater than 67 percent starting in the first year a CEO displays this
	behavior. Option moneyness is calculated as follows: first, we calculate the
	realizable value per option as the total realizable value of the exercisable options
	divided by the number of exercisable options [Value_Per_option =
	(OPT_UNEX_EXER_EST_VAL / OPT_UNEX_EXER_NUM)]. Second, we
	compute the estimate of the average exercise price of the options by subtracting
	the per-option realizable value from the stock price at the fiscal year-end
	[avg_exercise_price = (prccf - Value_Per_option)]. Lastly, the average percent
	moneyness of an option equals the per-option realizable value divided by the
	estimated average exercise price [avg_pctg_moneyness_opt = (Value_Per_option
	/ avg_exercise_price)]. [Malmendier and Tate, 2005; Campbell et al. 2011;
	Hirshleifer et al. 2012]
Net_buyer	An indicator variable that equals one if the number of years at which a CEO is a
	net-buyer is higher than those at which she/he is a net seller. Net_buyer is
	calculated as follows: first, we compute the net stock purchases by a CEO as
	purchases minus sales, both in units of shares [net_purchase = (
	SHROWN_EXCL_OPTS _t - SHROWN_EXCL_OPTS _{t-1} )], then we calculate the
	number of years at which a CEO has bought more shares than he/she sold.
	[Malmendier and Tate, 2005; Campbell et al. 2011]

# Table A1. Alternative measures of CEO political Ideology (Republican)

This table presents tests of the association between CEO political ideology and management earnings forecast using alternative measures of CEO Republican ideology (Panel A & B), measures of CEO Democratic ideology (Panel C & D) as well as Other ideologies (Panel E & F). All models include control variables, year, and industry fixed effects. All variables are defined in Appendix A. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

			Panel A	. Alternati	ve measur	es of Repu	blican ideo	ology (1)				
	Iss	sue	Freq	uency	Ra	nge	Ln(He	orizon)	Opt	Bias	Foreca	st_Miss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
<i>Rep_index</i> _{year}	$0.102^{\dagger}$		$0.080^{\dagger}$		0.109†		$0.087^{\dagger}$		-0.092†		-0.106†	
	(3.52)		(2.83)		(3.56)		(3.42)		(-2.88)		(-3.10)	
<i>Rep_index</i> _{tenure}		$0.142^{\dagger}$		$0.126^{\dagger}$		$0.141^{\dagger}$		$0.119^{\dagger}$		-0.134 [†]		-0.157†
		(4.72)		(4.23)		(4.45)		(4.49)		(-4.03)		(-4.43)
Pseudo / Adj. $R^2$	0.257	0.257	0.279	0.280	0.253	0.253	0.268	0.269	0.220	0.220	0.176	0.176
			Panel B	. Alternati	ve measure	es of Repu	blican ideo	ology (2)				
	Асси	ıracy	Bad	News	Good	News	Positive	Surprise	Negative	Surprise	Neutral	Surprise
<i>Rep_index</i> _{year}	$0.057^{\dagger}$		$0.072^{**}$		0.038		$0.058^{**}$		-0.067**		0.003	
	(3.40)		(2.38)		(1.17)		(2.35)		(-2.53)		(0.08)	
<i>Rep_index</i> _{tenure}		$0.082^{\dagger}$		0.126†		0.041		0.051**		$-0.076^{\dagger}$		0.032
		(4.65)		(4.01)		(1.19)		(1.98)		(-2.75)		(0.77)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951
Pseudo / Adj. R ²	0.266	0.266	0.257	0.257	0.180	0.180	0.052	0.052	0.102	0.102	0.077	0.077

# Table A2. Alternative measures of CEO political Ideology (Democratic & Other)

This table presents tests of the association between CEO political ideology and management earnings forecast using measures of CEO Democratic ideology (Panel A & B) as well as Other ideologies (Panel C & D). All models include control variables, year, and industry fixed effects. All variables are defined in Appendix A. T-statistics are computed using robust standard errors and reported in parentheses.  † ,  **  and  *  denote significance at the 1%, 5%, and 10% levels, respectively.

	, -	-,	10/010	/	anel A.	Alterna	tive me	asures o	f Demo	cratic ic	leology	(1)						
		Issue		F	requent	cy		Range		Lr	n(Horizo	on)	(	<b>OptBias</b>		Fo	recast N	Miss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)	(17)	(18)
Dem_dum _{cycle}	0.115			0.121			0.046			0.078			-0.197**			-0.193*		
	(1.26)			(1.33)			(0.47)			(0.94)			(-2.02)			(-1.88)		
Dem_dum _{cycle2}		0.007			-0.008			-0.082			-0.015			-0.022			-0.002	
		(0.12)			(-0.15)			(-1.40)			(-0.30)			(-0.37)			(-0.03)	
Dem_dum _{tenure}			-0.044			-0.007			-0.065			-0.059			0.018			0.136
			(-0.61)			(-0.10)			(-0.86)			(-0.89)			(0.22)			(1.59)
Pseudo/Adj.R ²	0.257	0.257	0.257	0.279	0.279	0.279	0.253	0.253	0.253	0.268	0.268	0.268	0.220	0.220	0.220	0.176	0.176	0.176
				Р	anel B.	Alternat	tive me	asures o	f Demo	cratic ic	leology	(2)						
	Ŀ	Accurac	y	В	ad_New	VS	G	ood_Ne	WS	Posit	tive_Sur	prise	Negat	tive_Sur	prise	Neu	tral_Sur	prise
Dem_dum _{cycle}	0.073			0.161*			0.020			-0.020			-0.030			0.079		
	(1.34)			(1.71)			(0.20)			(-0.25)			(-0.36)			(0.68)		
Dem_dum _{cycle2}		0.016			0.055			0.051			-0.050			0.020			0.053	
		(0.48)			(0.94)			(0.82)			(-1.05)			(0.40)			(0.73)	
Dem_dum _{tenure}			-0.022			-0.039			0.116			0.012			0.049			-0.156
			(-0.50)			(-0.51)			(1.49)			(0.20)			(0.74)			(-1.53)
Pseudo/Adj. R ²	0.258	0.258	0.258	0.257	0.257	0.257	0.180	0.180	0.180	0.052	0.052	0.052	0.102	0.102	0.102	0.077	0.077	0.077
					Panel	C. Alteri	native r	neasures	s of Oth	er ideol	ogies (1	)						
		Issue		I	Frequen	су		Range	2	L	n(Horiz	zon)		<b>OptBid</b>	is	Fa	precast_	Miss
Other_Index	0.043			0.265†			0.035			0.093			-0.027	1		0.03	1	
	(0.67)			(4.03)			(0.52)	)		(1.62)	)		(-0.39)	)		(0.41	)	
Other_dum _{cycle}		0.015			$0.188^{\dagger}$			0.042			0.056	,		-0.012	2		0.05	3
		(0.28)			(3.49)			(0.77)			(1.19)	)		(-0.22	)		(0.85	j)
All_dum _{cycle}			0.014			0.112**			0.070			0.039	)		-0.01	0		0.021
			(0.27)			(2.13)			(1.26)			(0.85)	)		(-0.17	')		(0.34)
Pseudo /Adj. R ²	0.257	0.257	0.257	0.280	0.279	0.279	0.253	0.253	0.253	0.268	0.268	0.268	0.220	0.220	0.220	0.17	6 0.17	6 0.176

#### Table A2. Alternative measures of CEO political Ideology (Democratic & Other). Cont'd

Panel D. Alternative measures of Other ideologies (2)												
	Accuracy	Bad_News	Good_News	Positive_Surprise	Negative_Surprise	Neutral_Surprise						
Other_Index	$0.064^{*}$	0.078	0.088	0.029	0.005	-0.061						

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	(1.70)			(1.18)			(1.23)			(0.53)			(0.08)			(-0.67)		
Other dum _{cycle}		0.038			0.055			0.048			0.005			-0.011			0.032	
		(1.24)			(1.00)			(0.81)			(0.11)			(-0.23)			(0.44)	
All_dum _{cycle}			0.005			-0.003			0.063			0.021			-0.030			-0.033
			(0.17)			(-0.05)			(1.08)			(0.47)			(-0.62)			(-0.46)
Pseudo /Adj. $R^2$	0.258	0.258	0.259	0.257	0.256	0.257	0.180	0.180	0.180	0.052	0.052	0.052	0.102	0.102	0.102	0.077	0.077	0.077
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951

#### Table A3. Alternative measures of CEO overconfidence

This table presents tests of the association between CEO political ideology and management earnings forecast using Net_buyer as an alternative measure of CEO overconfidence and controlling for CEO characteristics (Ln(Tenure), Ln(Age), Duality, CEO Gender, Ln(Delta), Ln(Vega), and CEO_Own, in addition to baseline control variables. All variables are defined in Appendix A. All models include year and industry fixed effects. T-statistics are computed using robust standard errors and reported in parentheses.[†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

			Panel A. C	EO Politica	l ideology a	and manage	ment earnir	igs forecast				
-	Iss	sue	Freq	uency	Ra	nge	Ln(Ho	rizon)	Opt	Bias	M	iss
_	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Rep_Dum	$0.122^{\dagger}$		0.183†		0.121†		$0.114^{\dagger}$		-0.135†		-0.124†	
	(3.27)		(5.04)		(3.11)		(3.45)		(-3.32)		(-2.87)	
Rep_Index		$0.108^{**}$		$0.124^{\dagger}$		$0.128^{\dagger}$		$0.094^{**}$		-0.143†		<b>-</b> 0.173 [†]
		(2.35)		(2.77)		(2.68)		(2.31)		(-2.86)		(-3.28)
Net_buyer	0.008	0.011	0.024	0.027	0.046	0.048	-0.005	-0.003	-0.035	-0.037	-0.044	-0.047
	(0.22)	(0.28)	(0.72)	(0.79)	(1.12)	(1.18)	(-0.15)	(-0.10)	(-0.79)	(-0.86)	(-0.96)	(-1.01)
CEO Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	25,562	25,562	25,562	25,562	25,562	25,562	25,562	25,562	25,562	25,562	25,562	25,562
<i>Pseudo</i> $R^2/Adj$ . $R^2$	0.250	0.249	0.289	0.289	0.248	0.248	0.273	0.273	0.216	0.216	0.174	0.174
			Panel B. C	EO Politica	l ideology a	and manage	ment earnin	igs forecast	(2)			
	Acu	racy	Bad	News	Good	News	Positive	Surprise	Negative	Surprise	Neutral	Surprise
	(13)	(14)	(15)	(16)	(17)	(18)	(19)	(20)	(21)	(22)	(23)	(24)
Rep Dum	0.080†		0.124†		0.042		0.058*		-0.087**		0.025	
* =	(3.69)		(3.20)		(1.03)		(1.82)		(-2.52)		(0.49)	
Rep Index	. ,	$0.075^{\dagger}$		$0.119^{**}$	, ,	-0.010		0.024		$-0.072^{*}$		0.084
× =		(2.78)		(2.49)		(-0.20)		(0.62)		(-1.71)		(1.35)
Net buyer	-0.002	-0.000	0.031	0.033	-0.058	-0.057	$0.061^{*}$	0.061*	-0.053	-0.054	-0.057	-0.056
	(-0.08)	(-0.02)	(0.75)	(0.81)	(-1.36)	(-1.34)	(1.87)	(1.89)	(-1.51)	(-1.55)	(-1.11)	(-1.09)
CEO Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	25,562	25,562	25,562	25,562	25,562	25,562	25,562	25,562	25,562	25,562	25,562	25,562
Pseudo $R^2/Adj$ . $R^2$	0.274	0.274	0.252	0.252	0.168	0.168	0.0525	0.0524	0.102	0.102	0.0779	0.0780

# Table A4. Subsamples based on CEO donation activity.

This table presents results using a restricted sample of firms in which CEOs make at least one donation during the sample period (Panels A & B) and a restricted sample of donation years (Panel C & D). All variables are defined in Appendix A. All models include control variables, year, and industry fixed effects. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively. n=21,042 in Panel A and 12,258 in Panel B

570, and 1070 level	-,				,	ctivity sub	sample (1)					
	Iss	sue	Freq	uency		nge	• • • • • • • • • • • • • • • • • • • •	orizon)	Opti	Bias	Foreca	st_Miss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Rep_Dum	0.099†		0.119†		0.093**		0.077**		-0.092**		-0.118†	
	(2.68)		(3.43)		(2.42)		(2.40)		(-2.30)		(-2.77)	
Rep_Index		$0.082^*$		0.046		$0.101^{**}$		0.053		-0.091*		-0.165†
		(1.84)		(1.10)		(2.14)		(1.37)		(-1.88)		(-3.23)
Pseudo / Adj. $R^2$	0.263	0.263	0.297	0.296	0.256	0.256	0.281	0.281	0.221	0.221	0.175	0.175
		ıracy		News	Good	News	Positive	Surprise	Negative	Surprise	Neutral	Surprise
Rep_Dum	$0.068^{\dagger}$		$0.085^{**}$		0.028		$0.061^{*}$		$-0.088^{\dagger}$		0.012	
	(3.25)		(2.21)		(0.70)		(1.96)		(-2.64)		(0.24)	
Rep_Index		$0.056^{**}$		$0.083^{*}$		-0.025		0.046		-0.084**		0.051
		(2.22)		(1.77)		(-0.50)		(1.22)		(-2.08)		(0.85)
Pseudo / Adj. $R^2$	0.279	0.279	0.265	0.265	0.175	0.175	0.045	0.045	0.096	0.095	0.083	0.083
				Panel B. I	Donation a	ctivity subs	sample (2)					
	Issi	ие	Freque	ncy	Range	2	Ln(Horiz	zon)	OptBia	S	Forecas	t_Miss
Rep_Dum	$0.161^{\dagger}$		$0.155^{\dagger}$		$0.156^{\dagger}$		$0.144^{\dagger}$		-0.141†		<b>-</b> 0.166 [†]	
	(3.32)		(3.20)		(3.05)		(3.39)		(-2.68)		(-3.01)	
Rep_Index		0.124**		0.021		0.150**		0.099**		-0.115*		-0.215†
2		(2.18)		(0.37)		(2.49)		(2.03)		(-1.87)		(-3.32)
Pseudo / Adj. R ²	0.272	0.272	0.305	0.304	0.270	0.270	0.291	0.291	0.231	0.231	0.183	0.184
		ıracy		News	Good	News		Surprise		Surprise	Neutral	Surprise
Rep_Dum	$0.109^{\dagger}$		0.134†		0.032		$0.084^{**}$		-0.101**		-0.023	
	(3.85)		(2.68)		(0.61)		(2.06)		(-2.27)		(-0.36)	
Rep_Index		$0.080^{**}$		$0.110^{*}$		-0.052		0.042		-0.077		0.045
2		(2.45)		(1.85)		(-0.84)		(0.87)		(-1.49)		(0.58)
<i>Pseudo / Adj.</i> $R^2$	0.292	0.292	0.272	0.272	0.170	0.170	0.045	0.045	0.097	0.097	0.102	0.102

#### Table A5. PSM. Alternative measures of CEO political ideology.

This table presents the test of the difference in management earnings forecast between firms with Republican CEOs and a sample of control firms with non-Republican CEOs matched primarily on firm characteristics, year, and industry—Panel A, B, and present results using  $Rep_dum_{only}$   $Rep_dum_{cycle}$ , and  $Rep_dum_{tenure}$ , respectively. In panel A, treatment denotes  $Rep_dum_{only}$ , which is an indicator variable that equals 1 if all donations of a CEO in an election cycle are directed to the Republican Party only (neither Democratic nor others) and control refers to a matching sample of CEOs who donated to other parties or never donated. In panel B, *treatment* denotes  $Rep_dum_{cycle}$ , which is an indicator variable that equals one if all donations of a CEO in an election cycle are directed to the Republican Party and control refers to matching sample if the donations of a CEO in an election cycle are directed to the Republican Party and control refers to matching sample if the donations of a CEO in an election cycle are directed to the Republican Party. In panel C, treatment denotes  $Rep_dum_{tenure}$ , which is an indicator variable that equals one if all donations of a CEO during her/his entire tenure are directed to the Republican Party and control refers to matching sample if all donations of a CEO during her/his entire tenure are directed to the Republican Party. All variables are defined in Appendix A. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

			Ра	nnel A. PSM. C	CEO ideolo	ogy is measu	red using Rep	o_dum _{only}				
	Issue	Frequency	Range	Ln(Horizon)	<b>OptBias</b>	Forecast_ Miss	Accuracy	Bad_ News	Good_ News	Positive_ Surprise	Negative_ Surprise	Neutral_ Surprise
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Rep dumonly	0.132	0.133*	0.190**	0.133*	-0.222**	-0.198**	0.122**	0.133	0.004	0.111	-0.177**	0.138
	(1.58)	(1.79)	(2.08)	(1.78)	(-2.35)	(-1.99)	(2.55)	(1.47)	(0.04)	(1.53)	(-2.25)	(1.16)
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	3,426	3,426	3,426	3,426	3,426	3,426	3,426	3,426	3,426	3,426	3,426	3,426
Pseudo / Adj. $R^2$	0.231	0.249	0.259	0.237	0.220	0.185	0.244	0.254	0.171	0.0719	0.131	0.111
			Pa	nel B. PSM. C	CEO ideolo	gy is measur	ed using Rep	o dum _{cycle}				
Rep dum _{cycle}	0.132*	$0.187^{**}$	$0.208^{**}$	0.156**	-0.059	-0.112	$0.090^{**}$	0.112	-0.008	0.131**	-0.120*	-0.061
	(1.67)	(2.41)	(2.55)	(2.32)	(-0.71)	(-1.27)	(1.98)	(1.40)	(-0.09)	(1.98)	(-1.66)	(-0.59)
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	4,172	4,172	4,172	4,172	4,172	4,172	4,172	4,172	4,172	4,172	4,172	4,172
Pseudo /Adj. R ²	0.298	0.307	0.271	0.325	0.244	0.190	0.312	0.285	0.179	0.0591	0.116	0.114
			Pa	nel C. PSM. C	EO ideolo	gy is measur	ed using Rep	<i>dum_{tenure}</i>				
Rep dum _{tenure}	0.199*	$0.210^{**}$	$0.196^{*}$	0.211**	0.007	-0.155	$0.188^{\dagger}$	0.111	0.031	0.043	-0.072	0.104
	(1.87)	(2.00)	(1.81)	(2.33)	(0.06)	(-1.33)	(3.05)	(1.04)	(0.28)	(0.49)	(-0.75)	(0.70)
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
<b>Observations</b>	2,422	2,422	2,422	2,422	2,422	2,422	2,422	2,422	2,422	2,422	2,422	2,422
$Pseudo / Adj. R^2$	0.302	0.302	0.279	0.320	0.242	0.191	0.310	0.275	0.186	0.07	0.134	0.109

#### Table A6. Management earnings forecasts around CEO turnover. A DID test

This table presents estimates from the Difference-in-Difference (DID) regressions of the association between CEO political ideology and management earnings forecasts around CEO turnover events (-2, +2). *After* is an indicator variable equals one for the years after the CEO turnover. We only consider turnover events where long-term old CEOs are replaced by long-term new CEOs (long-term is defined as holding the position for at least two years). *Rep_Leaving* is an indicator variable equals 1 if a firm replaces a Rep CEO with a non-Rep CEO, 0 otherwise. Republican CEOs are defined using *Rep_dum_{Only}*, which is an indicator variable that equals 1 if all donations of a CEO in an election cycle are directed to the Republican Party only (neither Democratic nor others). All models include control variables, year, and industry fixed effects. All control variables are defined in Appendix A. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

	Issue	Frequency	Range	Ln	<b>OptBias</b>	Forecast_	Accuracy	Bad_	$Good_{-}$	Positive_	Negative_	Neutral_
	Issue	Frequency	Kunge	(Horizon)	Opibius	Miss	Ассигису	News	News	Surprise	Surprise	Surprise
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
After*												
Rep Leaving	-0.215*	-0.012	-0.119	-0.150	$0.406^{\dagger}$	$0.432^{\dagger}$	-0.135*	-0.207*	0.010	-0.191*	0.261**	-0.149
	(-1.83)	(-0.11)	(-0.98)	(-1.35)	(3.21)	(3.27)	(-1.89)	(-1.71)	(0.08)	(-1.90)	(2.41)	(-0.90)
Rep_Leaving	0.109	0.017	-0.033	0.071	-0.199*	-0.285 [†]	$0.107^{*}$	0.109	-0.051	-0.000	-0.021	0.049
	(1.15)	(0.20)	(-0.33)	(0.81)	(-1.95)	(-2.70)	(1.88)	(1.12)	(-0.48)	(-0.00)	(-0.24)	(0.38)
After	$0.109^{*}$	0.130**	$0.184^{\dagger}$	0.085	-0.250†	-0.216 [†]	0.040	$0.174^{+}$	-0.050	-0.033	0.047	-0.020
	(1.87)	(2.50)	(2.85)	(1.58)	(-3.74)	(-3.05)	(1.15)	(2.78)	(-0.77)	(-0.65)	(0.87)	(-0.24)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo / Adj. $R^2$	0.224	0.292	0.228	0.256	0.205	0.155	0.270	0.232	0.151	0.054	0.101	0.073
Observations	11,815	11,815	11,815	11,815	11,815	11,815	11,815	11,815	11,815	11,815	11,815	11,815

# Table A7. The effect of change in CEO political ideology due to CEO turnover on Change in management earnings forecasts.

This table presents tests of the association between changes in CEO political ideology due to CEO turnover and changes in management earnings forecasts.  $\Delta dependent$  is the difference between the first full fiscal year under the new CEO and the last full fiscal year under the old CEO.  $\Delta REP_{CEO}$  is defined as the changes in CEO political ideology due to CEO turnover, where  $\Delta REP_{CEO} = 1$  if a Republican CEO ( $Rep_dum_{Only}$ ), 0 if the political ideology is similar after a CEO turnover, and -1 if a Democratic CEO replaces a Republican minded CEO. Panel A reports results for all CEO turnover events. Panel B reports results only when an old CEO is in position for at least three years. All models include control variables, year, and industry fixed effects. All variables are defined in Appendix A. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

				Par	nel A. CEO	turnover	sample					
		∆Frequenc		$\Delta Ln$		∆Forecas	rt -				∆Negative	2
	∆Issue	21 ^r requenc	⊿Range	(Horizon)	$\Delta OptBias$	_	∆Accuracy	$\Delta Bad_{-}$	$\Delta Good_{-}$	$\Delta Positive_$	_	$\Delta Neutral_$
		y		(110/12011)		Miss		News	News	Surprise	Surprise	Surprise
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
$\Delta REP_{CEO}$	0.201**	0.199	$0.247^{**}$	$0.883^{*}$	-0.307**	-0.365†	$0.782^{**}$	0.193	0.002	-0.059	0.012	0.047
	(2.28)	(0.49)	(2.36)	(1.88)	(-2.26)	(-3.01)	(2.26)	(1.59)	(0.02)	(-0.37)	(0.08)	(0.49)
$\Delta Controls$	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	172	172	172	172	172	172	172	172	172	172	172	172
Pseudo/Adj. $R^2$	0.489	0.483	0.463	0.458	0.437	0.474	0.418	0.438	0.369	0.451	0.431	0.533
				Panel B. L	ong-term o	ld CEO t	urnover samp	ole				
		1 Englishong		$\Delta Ln$		∆Forecas	t				<b>ANegative</b>	2
	∆Issue	∆Frequenc	⊿Range	(Horizon)	∆ <i>OptBias</i>	_	$\Delta Accuracy$	$\Delta Bad_{-}$	$\Delta Good_{-}$	$\Delta Positive_$	_	$\Delta Neutral_$
		y		(110/12011)		Miss		News	News	Surprise	Surprise	Surprise
$\varDelta REP_{CEO}$	$0.158^{*}$	-0.228	$0.274^{**}$	0.633	$-0.240^{*}$	-0.413†	$0.719^{*}$	0.174	-0.104	-0.004	-0.049	0.054
	(1.79)	(-0.56)	(2.48)	(1.39)	(-1.75)	(-2.95)	(1.76)	(1.43)	(-0.98)	(-0.03)	(-0.28)	(0.46)
<b>∆Controls</b>	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
<b>Observations</b>	162	162	162	162	162	162	162	162	162	162	162	162
Pseudo/Adj. $R^2$	0.534	0.504	0.476	0.521	0.469	0.493	0.421	0.457	0.404	0.500	0.485	0.547

# Table A8. Controlling for CEO turnover and tenure.

This table presents results when excluding CEO turnover years (Panels A & B), and the first three years of CEO tenure (Panels C & D). All models include control variables, year, and industry fixed effects. All variables are defined in Appendix A. T-statistics are computed using robust standard errors and reported in parentheses.  † , ** and * denote significance at the 1%, 5%, and 10% levels, respectively. n = 30,319 in Panels A & B and 20,681 in Panels C & D.

*	, ,		U		,	Panel A	. Exclu	ding CE	EO turno	over yea	rs (1)							
		Issue		F	requenc	су		Range		Lr	n(Horizo	on)		<b>OptBias</b>	5	For	recast_l	Miss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)	(17)	(18)
Rep_Dum	$0.140^{\dagger}$			0.166†			0.133 [†]			$0.118^{\dagger}$			-0.152*			-0.149†		
	(4.02)			(5.07)			(3.63)			(3.94)			(-3.98)			(-3.66)		
Rep_Index		0.139†			0.121 [†]			0.161 [†]			$0.106^{\dagger}$			$-0.178^{\dagger}$			-0.205†	•
		(3.22)			(3.01)			(3.54)			(2.88)			(-3.77)			(-4.11)	
<i>Rep_index</i> _{year}			$0.111^{\dagger}$			$0.090^{\dagger}$			0.115†			$0.095^{\dagger}$			-0.099†			-0.110 [†]
			(3.68)			(3.05)			(3.61)			(3.60)			(-2.98)			(-3.10)
$Pseudo / Adj.R^2$	0.260	0.260	0.260	0.282	0.281	0.281	0.257	0.257	0.257	0.271	0.271	0.271	0.222	0.222	0.221	0.179	0.179	0.178
						Panel B	B. Exclu	ding CE	EO turno	over yea	rs (2)							
	I	Accurac	y	В	ad_Nev	VS	G	ood_Ne	WS	Posi	tive_Sur	prise	Nega	tive_Su	rprise	Neut	tral_Sur	rprise
Rep_Dum	$0.088^{\dagger}$			$0.150^{+}$			0.052			$0.064^{**}$			-0.097†			0.028		
	(4.44)			(4.14)			(1.33)			(2.16)			(-3.06)			(0.59)		
Rep_Index		$0.088^{\dagger}$			0.169†			0.005			0.045			-0.095**	k		0.080	
		(3.59)			(3.74)			(0.10)			(1.23)			(-2.43)			(1.38)	
<i>Rep_index</i> _{year}			$0.064^{\dagger}$			$0.083^{\dagger}$			0.046			0.056**			-0.071**	k		0.014
			(3.66)			(2.61)			(1.33)			(2.18)			(-2.56)			(0.35)
<i>Pseudo/Adj.</i> $R^2$	0.266	0.266	0.266	0.259	0.259	0.259	0.173	0.173	0.173	0.050	0.050	0.050	0.099	0.099	0.099	0.077	0.077	0.077
					Par	nel C. Ez	xcludin	g first 3	years of	f CEO to	enure (1	)						
		Issue			requent			Range			n(Horize	on)		OptBia.	5		recast_1	Miss
Rep_Dum	$0.107^{\dagger}$			$0.147^{\dagger}$			0.102**			0.095†			-0.124	İ		-0.109*	*	
	(2.61)			(3.83)			(2.36)			(2.67)			(-2.74)			(-2.26)		
Rep_Index		$0.101^{*}$			$0.094^{*}$			0.131**			$0.079^{*}$			-0.144*	*		$-0.170^{\circ}$	ŕ
		(1.94)			(1.96)			(2.38)			(1.78)			(-2.53)			(-2.83)	
<i>Rep_index</i> _{year}			$0.095^{\dagger}$			$0.077^{**}$			0.112 [†]			$0.079^{\dagger}$			-0.094*	*		-0.097**
			(2.72)			(2.28)			(3.02)			(2.58)			(-2.43)			(-2.37)
Pseud/Adj. $R^2$	0.262	0.262	0.262	0.283	0.282	0.282	0.258	0.258	0.259	0.275	0.275	0.275	0.221	0.221	0.221	0.180	0.180	0.180

Tuble Hor Colle	i oning i	U CL	, tui 110	or una	venar e	00110	•											
					Par	nel D. Ez	xcluding	g first 3	years of	f CEO te	enure (2	)						
	I	Accurac	y	В	ad New	VS	G	ood Ne	ws	Posit	tive Sur	prise	Nega	tive Su	rprise	Neut	ral Sur	prise
Rep Dum	$0.067^{\dagger}$			0.101**			0.039			0.102 [†]		-	-0.122 [†]		-	-0.013		
	(2.89)			(2.35)			(0.84)			(2.95)			(-3.24)			(-0.24)		
Rep Index		0.063**			0.112**			-0.018			$0.076^{*}$		. ,	-0.121 [†]		. ,	0.058	
		(2.18)			(2.05)			(-0.29)			(1.75)			(-2.57)			(0.85)	
Rep index _{year}			0.050**			$0.069^{*}$		. ,	0.025		. ,	0.075**		. ,	$-0.092^{\dagger}$		. ,	0.013
			(2.47)			(1.90)			(0.62)			(2.52)			(-2.87)			(0.27)
Pseudo/Adj. $R^2$	0.270	0.270	0.270	0.260	0.260	0.260	0.179	0.179	0.179	0.050	0.050	0.050	0.098	0.098	0.098	0.082	0.082	0.082
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes

# Table A8. Controlling for CEO turnover and tenure. Cont'd

#### Table A9. Change-on-change regression

This table presents tests of the association between CEO political ideology and management earnings forecast where all dependent and independent variables are annual changes. We exclude the firm-years with 0 changes in either dependent or independent variables. All models include control variables, firm, and year fixed effects. All control variables are defined in Appendix A. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

				Panel A.	Change-or	n-change r	models (1)					
	∆Is	sue	∆Freq	juency	∆Rc	inge	∆Ln(H	orizon)	∆Ор	tBias	∆Forec	ast_Miss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
$\Delta Rep_index_{year}$	$0.168^{\dagger}$		0.036		0.059		$0.152^{\dagger}$		-0.045		-0.025	
	(2.80)		(0.60)		(0.89)		(3.29)		(-1.24)		(-0.70)	
$\Delta Rep_dum_{Only}$		$0.382^{*}$		$0.300^{*}$		0.292		0.399†		-0.282**		-0.270**
		(1.73)		(1.94)		(0.67)		(3.63)		(-2.43)		(-2.15)
<i>∆controls</i>	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm & Year FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	1,118	315	3,203	794	1,013	266	3,857	973	1,594	409	1,594	393
Pseudo/Adj. $R^2$	0.518	0.814	0.240	0.523	0.459	0.846	0.298	0.535	0.414	0.777	0.373	0.723
				Panel B.	Change-or	n-change r	models (2)					
	AAcc	uracy	ARad	News	AGood	l News	∆Positive	e_Surpris	∆Negativ	ve_Surpri	∆Neutra	l_Surpris
		ur de y	-					е		ie –		e
$\Delta Rep_index_{year}$	0.126†		0.050		$0.061^{*}$		$0.047^{**}$		-0.055**		0.010	
	(3.46)		(1.04)		(1.81)		(1.99)		(-2.15)		(0.22)	
$\Delta Rep_dum_{Only}$		$0.335^{\dagger}$		-0.040		0.165		-0.028		0.039		0.308
		(3.49)		(-0.19)		(1.28)		(-0.40)		(0.50)		(1.63)
<i>∆controls</i>	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm & Year FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	3,600	905	1,256	334	1,699	403	3,590	924	3,078	812	1,360	340
Pseudo/Adj. $R^2$	0.252	0.498	0.437	0.795	0.429	0.758	0.281	0.534	0.327	0.578	0.368	0.779

#### Table A10. Controlling for CEO characteristics, incentives, and managerial ability

This table presents results of tests that control for managerial ability, *MA_Score*, controlling for CEO characteristics (Ln(Tenure), Ln(Age), Duality, Ln(Delta), Ln(Vega), CEO_Own, and managerial ability) in addition to the baseline control variables. All variables are defined in Appendix A. All models include year and industry fixed effects. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

	e sigiii				I	Panel A.	Control	ling for	manage	rial abil	ity (1)							
		Issue		F	requenc	<i>y</i>		Range		Li	n(Horiza	on)		<b>OptBias</b>		For	ecast N	liss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)	(17)	(18)
Rep_Dum	0.121†			$0.173^{\dagger}$			0.121†			$0.114^{\dagger}$			-0.128†			-0.122 [†]		
	(3.44)			(5.17)			(3.27)			(3.75)			(-3.34)			(-2.99)		
Rep_Index		0.102**			$0.111^{\dagger}$			0.116**			0.090**			-0.123†			-0.161*	
		(2.36)			(2.73)			(2.57)			(2.39)			(-2.60)			(-3.23)	
<i>Rep_index</i> _{year}			$0.109^{\dagger}$			$0.089^{\dagger}$			$0.107^{\dagger}$			$0.100^{\dagger}$			-0.094†			-0.105†
			(3.60)			(2.99)			(3.35)			(3.76)			(-2.83)			(-2.97)
MA_Score				-0.361†										0.538†	0.532†		0.583†	$0.576^{\dagger}$
	(-4.63)	(-4.65)	(-4.62)	(-2.94)	(-2.98)	(-2.93)	(-5.08)	(-5.11)	(-5.08)	(-6.01)	(-6.04)	(-6.00)	(3.95)	(3.97)	(3.93)	(4.04)	(4.06)	(4.02)
Observations	30,638	30,638	30,638	30,638	30,638	30,638	)	)	30,638	30,638	30,638	30,638	30,638	30,638	30,638	30,638	30,638	30,638
<i>Pseudo / Adj.</i> $R^2$	0.262	0.262	0.262	0.290	0.290				0.255		0.277	0.277	0.223	0.223	0.223	0.178	0.178	0.178
							Control	ling for	manage	rial abili	ity (2)							
		Accuracy	у		ad_New	'S	G	ood_Nei	VS	Posi	tive_Sur			tive_Sur	rprise	Neut	ral_Sur	prise
Rep_Dum	$0.079^{\dagger}$			$0.112^{\dagger}$			0.061			0.043			-0.071**			0.027		
	(3.94)			(3.07)			(1.56)			(1.42)			(-2.21)			(0.58)		
Rep_Index		$0.064^{\dagger}$			$0.100^{**}$			0.013						0.062				
		$0.064^{\dagger}$									0.021			-0.062			0.072	
		(2.61)			(2.20)			(0.013)			(0.58)			(-1.57)			(1.24)	
Rep_indexyear			$0.060^{\dagger}$		(2.20)	0.071**			0.049		(0.58)	0.053**		(-1.57)	-0.069**			0.021
		(2.61)	(3.41)		(2.20)	(2.25)		(0.27)	(1.44)		(0.58)	(2.06)		(-1.57)	(-2.46)		(1.24)	(0.53)
Rep_index _{year} MA_Score		(2.61) -0.483 [†]	(3.41) -0.480 [†]	-0.531†	(2.20) -0.534 [†]	(2.25) -0.530 [†]		(0.27) -0.319**	(1.44) -0.317**		(0.58) -0.292 [†]	(2.06) -0.290 [†]	0.271**	(-1.57) 0.272**	(-2.46) 0.270**	-0.139	(1.24)	(0.53) -0.137
MA_Score	(-6.69)	(2.61) -0.483 [†] (-6.71)	(3.41) -0.480 [†] (-6.68)	-0.531 [†] (-4.02)	(2.20) -0.534 [†] (-4.04)	(2.25) -0.530 [†] (-4.01)	(-2.27)	(0.27) -0.319** (-2.29)	(1.44) -0.317** (-2.27)	(-2.71)	(0.58) -0.292 [†] (-2.72)	(2.06) -0.290 [†] (-2.70)	0.271 ^{**} (2.33)	(-1.57) 0.272** (2.34)	(-2.46) 0.270** (2.32)	-0.139 (-0.85)	(1.24) -0.140 (-0.86)	(0.53) -0.137 (-0.84)
MA_Score CEO controls	(-6.69) Yes	(2.61) -0.483 [†] (-6.71) Yes	(3.41) -0.480 [†] (-6.68) Yes	-0.531 [†] (-4.02) Yes	(2.20) -0.534 [†] (-4.04) Yes	(2.25) -0.530 [†] (-4.01) Yes	(-2.27) Yes	(0.27) -0.319** (-2.29) Yes	(1.44) -0.317** (-2.27) Yes	(-2.71) Yes	(0.58) -0.292 [†] (-2.72) Yes	(2.06) -0.290 [†] (-2.70) Yes	0.271** (2.33) Yes	(-1.57) 0.272** (2.34) Yes	(-2.46) 0.270** (2.32) Yes	-0.139 (-0.85) Yes	(1.24) -0.140 (-0.86) Yes	(0.53) -0.137 (-0.84) Yes
MA_Score CEO controls Controls	(-6.69) Yes Yes	(2.61) -0.483 [†] (-6.71) Yes Yes	(3.41) -0.480 [†] (-6.68) Yes Yes	-0.531 [†] (-4.02) Yes Yes	(2.20) -0.534 [†] (-4.04) Yes Yes	(2.25) -0.530 [†] (-4.01) Yes Yes	(-2.27) Yes Yes	(0.27) -0.319 ^{**} (-2.29) Yes Yes	(1.44) -0.317** (-2.27) Yes Yes	(-2.71) Yes Yes	(0.58) -0.292 [†] (-2.72) Yes Yes	(2.06) -0.290 [†] (-2.70) Yes Yes	0.271 ^{**} (2.33) Yes Yes	(-1.57) 0.272** (2.34) Yes Yes	(-2.46) 0.270** (2.32) Yes Yes	-0.139 (-0.85) Yes Yes	(1.24) -0.140 (-0.86) Yes Yes	(0.53) -0.137 (-0.84) Yes Yes
MA_Score CEO controls Controls Year & Ind. FE	(-6.69) Yes Yes Yes	(2.61) -0.483 [†] (-6.71) Yes Yes Yes	(3.41) -0.480 [†] (-6.68) Yes Yes Yes	-0.531 [†] (-4.02) Yes Yes Yes	(2.20) -0.534 [†] (-4.04) Yes Yes Yes	(2.25) -0.530 [†] (-4.01) Yes Yes Yes	(-2.27) Yes Yes Yes	(0.27) -0.319** (-2.29) Yes Yes Yes	(1.44) -0.317** (-2.27) Yes Yes Yes	(-2.71) Yes Yes Yes	(0.58) -0.292 [†] (-2.72) Yes Yes Yes	(2.06) -0.290 [†] (-2.70) Yes Yes Yes	0.271 ^{**} (2.33) Yes Yes Yes	(-1.57) 0.272** (2.34) Yes Yes Yes	(-2.46) 0.270** (2.32) Yes Yes Yes	-0.139 (-0.85) Yes Yes Yes	(1.24) -0.140 (-0.86) Yes Yes Yes	(0.53) -0.137 (-0.84) Yes Yes Yes
MA_Score CEO controls Controls	(-6.69) Yes Yes Yes 30,638	(2.61) -0.483 [†] (-6.71) Yes Yes Yes	(3.41) -0.480 [†] (-6.68) Yes Yes Yes	-0.531 [†] (-4.02) Yes Yes Yes	(2.20) -0.534 [†] (-4.04) Yes Yes Yes	(2.25) -0.530 [†] (-4.01) Yes Yes Yes	(-2.27) Yes Yes Yes	(0.27) -0.319 ^{**} (-2.29) Yes Yes	(1.44) -0.317** (-2.27) Yes Yes Yes	(-2.71) Yes Yes Yes	(0.58) -0.292 [†] (-2.72) Yes Yes Yes	(2.06) -0.290 [†] (-2.70) Yes Yes Yes	0.271 ^{**} (2.33) Yes Yes Yes	(-1.57) 0.272** (2.34) Yes Yes Yes	(-2.46) 0.270** (2.32) Yes Yes	-0.139 (-0.85) Yes Yes Yes	(1.24) -0.140 (-0.86) Yes Yes Yes	(0.53) -0.137 (-0.84) Yes Yes Yes

#### Table A11. Additional statistical specifications.

This table presents results using state fixed effects (Panel A & B), and standard errors clustered at the firm level (Panel C & D). All models include control variables, year, and industry fixed effects. All variables are defined in Appendix A. T-statistics are computed using robust standard errors and reported in parentheses.  † ,  **  and  *  denote significance at the 1%, 5%, and 10% levels, respectively.

•							Panel A	A. State	fixed eff	ects (1)			-					
		Issue		I	Frequent	cy		Range		Li	n(Horize	on)		<b>OptBias</b>	5	For	recast_N	Aiss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)	(17)	(18)
Rep_Dum	0.079**			$0.160^{\dagger}$			0.059			$0.078^{\dagger}$			-0.093*	*		-0.084**	•	
	(2.28)			(5.04)			(1.61)			(2.67)			(-2.45)			(-2.08)		
Rep_Index		0.058			$0.098^{**}$			0.047			0.053			-0.082*			-0.119**	ŧ
		(1.35)			(2.52)			(1.03)			(1.46)			(-1.73)			(-2.38)	
Rep_indexyear			$0.068^{**}$			0.069**			$0.056^{*}$			0.063**			-0.055*			-0.069**
			(2.27)			(2.45)			(1.79)			(2.46)			(-1.68)			(-1.96)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo/Adj. $R^2$	0.269	0.269	0.269	0.289	0.288	0.288	0.266	0.266	0.266	0.278	0.278	0.278	0.228	0.228	0.228	0.183	0.183	0.183
Observations	33,316	33,316	33,316	33,348	33,348	33,348	33,309	33,309	33,309	33,348	33,348	33,348	33,251	33,251	33,251	33,255	33,255	33,255
							Panel E	3. State f	ixed eff	ects (2)								
	Ē	1ccurac	v	l	Bad Nev	VS	Go	od New	s i	Positive	Surpris	е	Negativ	e Surpr	ise	Neutr	al Surp	orise
Rep_Dum	$0.067^{\dagger}$			$0.080^{**}$			0.026			0.058**			-0.084 [†]			0.012		
	(3.53)			(2.23)			(0.67)			(1.99)			(-2.69)			(0.26)		
Rep_Index		0.056**			0.066			-0.025			0.045			-0.080**			0.044	
		(2.38)			(1.47)			(-0.53)			(1.26)			(-2.10)			(0.76)	
Rep index _{year}		· · · ·	0.042**			0.033		· · · ·	0.016			$0.060^{**}$			-0.069**			-0.002
			(2.46)			(1.06)			(0.47)			(2.36)			(-2.55)			(-0.06)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo/Adj. R ²	0.277	0.277	0.277	0.268	0.268	0.268	0.188	0.188	0.188	0.054	0.054	0.054	0.104	0.104	0.104	0.082	0.082	0.082
Observations	33,348	33,348	33,348	33 <u>,29</u> 7	33,297	<u>33,29</u> 7	33,320	33,320	<u>33,32</u> 0	33,348	33,348	33,348	33,348	33,348	33,348	33,291	33,291	33,291

Table All. Aud	itional s	tatistica	i speem	ications														
					Pane	C. Stan	dard er	rors clus	tered at	the firm	n level (	1)						
		Issue		Fre	equency		Ì	Range		Ln	(Horizoi	n)	(	<b>OptBias</b>		For	ecast N	<i>Aiss</i>
Rep_Dum	0.128**	*		0.165**			0.127*			0.111**	k		-0.143*	*		-0.142*	k	
	(2.07)			(2.58)			(1.95)			(2.04)			(-2.53)			(-2.49)		
Rep_Index		0.126*			0.117			$0.144^{*}$			0.100			-0.158*	*		-0.194	
		(1.74)			(1.55)			(1.86)			(1.56)			(-2.35)	)		(-2.88)	
Rep_indexyear			0.102**	*		$0.080^*$			0.109**			$0.087^{**}$	•		-0.092**	*		-0.106**
			(2.37)			(1.73)			(2.33)			(2.26)			(-2.18)			(-2.47)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo / Adj. $R^2$	0.257	0.257	0.257	0.279	0.280	0.279	0.253	0.253	0.253	0.268	0.269	0.268	0.220	0.220	0.220	0.176	0.176	0.176
Observations	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951
					Pane	l D. Star	ndard er	rors clus	stered at	the firn	n level (2	2)						
		Асси			Bad N	lews		Good 1	Vews	Pos	sitive Sı	ırprise	Neg	ative Sı	urprise	Neu	tral Sur	prise
Rep_Dum	$0.087^{**}$			0.133**			0.052			$0.058^{*}$			-0.090†			0.029		
	(2.52)			(2.15)			(0.96)			(1.87)			(-2.75)			(0.58)		
Rep_Index		$0.084^{**}$			$0.144^{**}$			0.011			0.052			-0.094**	k		0.065	
		(2.07)			(1.98)			(0.16)			(1.37)			(-2.36)			(1.04)	
Rep_indexyear			$0.057^{**}$			$0.072^{*}$			0.038			0.058**			-0.067**			0.003
		(2.21)			(1.67)			(0.93	3)		(2.	28)		(-	-2.48)			(0.07)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
$Pseudo / Adj. R^2$	0.266	0.266	0.266	0.257	0.257	0.257	0.180	0.180	0.180	0.052	0.052	0.052	0.102	0.102	0.102	0.077	0.077	0.077
Observations	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951

#### Table A11. Additional statistical specifications. Cont'd

# Table A12. Political ideology vs. Political activism.

This table presents tests that attempt to differentiate between the political ideology and the political activism explanation of our baseline results. Panel A reports results for *Issue, Frequency, Range, Ln(Horizon), OptBias, and Forecast_Miss,* on the other hand. Panel B reports results for the models of the association between CEO political ideology on one hand and *Accuracy, Bad_News, Good_News, Positive_Surprise, Negative_Surprise, and Neutral_Surprise* on the other hand. All models include control variables, year, and industry fixed effects. All other independent variables are defined in Appendix A. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

				Panel A	A. Ideolog	y vs. activ	ism (1)					
	Iss	sue	Freq	uency	Ra	nge	Ln(Ho	orizon)	Opt	Bias	M	iss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Rep_dum	$0.117^{\dagger}$		$0.174^{\dagger}$		$0.120^{\dagger}$		$0.105^{\dagger}$		-0.138†		-0.130 [†]	
	(3.41)		(5.47)		(3.31)		(3.55)		(-3.63)		(-3.22)	
Dem_dum	-0.053		0.044		-0.031		-0.031		0.024		0.049	
	(-1.26)		(1.07)		(-0.71)		(-0.83)		(0.52)		(1.00)	
Rep_dum _{tenure}		$0.120^{\dagger}$		$0.110^{\dagger}$		$0.106^{**}$		$0.085^{**}$		-0.097**		-0.136†
		(2.72)		(2.65)		(2.28)		(2.23)		(-1.98)		(-2.64)
Dem_dum _{tenure}		-0.032		0.004		-0.054		-0.050		0.008		0.121
		(-0.44)		(0.06)		(-0.72)		(-0.76)		(0.10)		(1.42)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Industry FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951
Pseudo /Adj. $R^2$	0.257	0.257	0.280	0.279	0.253	0.253	0.269	0.268	0.220	0.220	0.176	0.176
				Panel 1	B. Ideolog	y vs. activ	ism (2)					
	Асси	uracy	Bad_	News	Good	News	Positive	Surprise	Negative	_Surprise	Neutral	Surprise
Rep_dum	$0.083^{\dagger}$		0.130†		0.052		0.046		$-0.080^{\dagger}$		0.038	
	(4.29)		(3.61)		(1.36)		(1.58)		(-2.58)		(0.81)	
Dem dum	-0.018		-0.016		0.001		-0.063*		0.051		0.043	
_	(-0.72)		(-0.36)		(0.01)		(-1.75)		(1.32)		(0.76)	
Rep dum _{tenure}		$0.069^{\dagger}$		$0.127^{\dagger}$		0.062		0.040		-0.066*		0.052
		(2.79)		(2.74)		(1.25)		(1.06)		(-1.66)		(0.87)
Dem dum _{tenure}		-0.014		-0.026		0.122		0.017		0.042		-0.150
—		(-0.33)		(-0.34)		(1.57)		(0.27)		(0.64)		(-1.47)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Industry FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951
Pseudo /Adj. $R^2$	0.266	0.266	0.257	0.257	0.180	0.180	0.052	0.052	0.102	0.102	0.077	0.077

#### Table A13. Robustness check. Active earnings forecast subsample

This table presents tests of the association between CEO political ideology and management earnings forecast using a subsample of the firm that have at least one earnings forecast during our sample period. All models include year and industry fixed effects. All other independent variables are defined in Appendix A. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

i			Pa	nel A. Act	ive earning	gs forecast	subsample	e (1)				
	Issu	ie	Frequer	ісу	Rang	re	Ln(Hor	rizon)	OptE	Bias	Foreca	st_Miss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Rep_Dum	0.123†		$0.180^{\dagger}$		0.122†		$0.110^{\dagger}$		-0.137†		-0.136†	
	(3.39)		(4.66)		(3.30)		(3.18)		(-3.60)		(-3.40)	
Rep_Index		0.126†		0.139†		$0.150^{+}$		$0.108^{**}$		-0.157†		-0.192†
		(2.85)		(2.96)		(3.31)		(2.56)		(-3.38)		(-3.95)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	25,437	25,437	25,437	25,437	25,437	25,437	25,437	25,437	25,437	25,437	25,437	25,437
Pseudo / Adj. $R^2$	0.222	0.222	0.295	0.294	0.226	0.226	0.261	0.261	0.192	0.192	0.149	0.149
			Pa	nel B. Act	ive earning	gs forecast	subsample	e (2)				
	Accur	acy	Bad_Ne	WS	$Good_N$	lews	Positive_S	Surprise	Negative_	Surprise	Neutral	Surprise
Rep_Dum	$0.088^{\dagger}$		0.125 [†]		0.046		$0.070^{**}$		-0.105†		0.020	
	(3.82)		(3.39)		(1.19)		(2.17)		(-2.97)		(0.41)	
Rep_Index		0.093†		$0.148^{\dagger}$		0.005		$0.072^{*}$		-0.107**		0.024
		(3.29)		(3.26)		(0.10)		(1.81)		(-2.48)		(0.40)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	25,437	25,437	25,437	25,437	25,437	25,437	25,437	25,437	25,437	25,437	25,437	25,437
Pseudo / Adj. $R^2$	0.269	0.269	0.226	0.226	0.144	0.144	0.043	0.043	0.090	0.089	0.076	0.076

#### Table A14. Cross-sectional test: High vs. low analyst coverage.

This table presents results for firms with high (above-median) analyst coverage (Panel A) and firms with low (below-median) analyst coverage (Panel B). All variables are defined in Appendix A. All models include year and industry fixed effects. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively

				Panel A	A1. High a	nalyst cove	rage (1)					
	Iss	ие	Freq	uency	Ra	nge	Ln(He	orizon)	Ор	tBias	Forec	ast_Miss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Rep_Dum	$0.147^{\dagger}$		$0.215^{\dagger}$		0.109**		$0.117^{\dagger}$		-0.164†		-0.165†	
	(3.20)		(4.71)		(2.28)		(3.04)		(-3.43)		(-3.28)	
Rep_Index		$0.107^{*}$		$0.106^{*}$		0.096		0.061		-0.145**		-0.223†
		(1.87)		(1.86)		(1.62)		(1.29)		(-2.46)		(-3.63)
Observations	16,003	16,003	16,006	16,006	16,006	16,006	16,006	16,006	16,003	16,003	16,003	16,003
Pseudo /Adj. $R^2$	0.291	0.290	0.325	0.324	0.285	0.285	0.319	0.319	0.235	0.235	0.186	0.186
				Panel A	A2. High a	nalyst cove	rage (2)					
	Асси	ıracy	Bad	News	Good	News	Positive	Surprise	Negative	Surprise	Neutral	Surprise
Rep_Dum	$0.094^{\dagger}$		0.135 [†]		0.031		0.027		-0.054		0.018	
	(3.52)		(2.85)		(0.64)		(0.72)		(-1.28)		(0.32)	
Rep_Index		$0.061^{*}$		$0.105^{*}$		-0.065		-0.013		-0.031		0.065
		(1.85)		(1.79)		(-1.06)		(-0.27)		(-0.60)		(0.91)
Observations	16,006	16,006	16,003	16,003	15,955	15,955	16,002	16,002	16,002	16,002	16,002	16,002
Pseudo /Adj. $R^2$	0.305	0.305	0.286	0.286	0.179	0.179	0.032	0.032	0.072	0.072	0.079	0.079
				Panel l	B1. Low ar	alyst cove	rage (1)					
	Iss	ие	Freq	uency	Ra	nge		orizon)	Opt	Bias	Foreca	st_Miss
Rep Dum	0.126**		$0.094^{**}$		$0.185^{\dagger}$		$0.099^{**}$		-0.131**		-0.124*	
	(2.50)		(2.33)		(3.42)		(2.33)		(-2.22)		(-1.96)	
Rep_Index	. ,	0.166†		$0.100^{**}$	. ,	$0.226^{\dagger}$	. ,	0.124**	. ,	-0.174**		-0.154**
_		(2.71)		(2.05)		(3.45)		(2.40)		(-2.42)		(-2.00)
Observations	17,945	17,945	17,945	17,945	17,945	17,945	17,945	17,945	17,898	17,898	17,877	17,877
Pseudo /Adj. $R^2$	0.226	0.227	0.227	0.227	0.239	0.239	0.224	0.224	0.206	0.206	0.172	0.172

				Panel I	B2. Low an	alyst cove	rage (2)					
	Ассі	ıracy	Bad	News	Good	News	Positive	Surprise	Negative	Surprise	Neutral	Surprise
Rep_Dum	$0.076^{+}$		0.144†		0.103*		0.125†		-0.147†		0.044	
	(2.91)		(2.67)		(1.75)		(2.90)		(-3.29)		(0.56)	
Rep Index		$0.098^{\dagger}$		$0.198^{\dagger}$		0.133*		0.130**		-0.162 [†]		0.075
		(3.09)		(3.01)		(1.85)		(2.50)		(-2.99)		(0.81)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	17,945	17,945	17,898	17,898	17,945	17,945	17,945	17,945	17,945	17,945	17,945	17,945
Pseudo $R^2/Adj$ . $R^2$	0.209	0.209	0.224	0.225	0.184	0.184	0.085	0.085	0.123	0.123	0.064	0.064

Table A14. Cross-sectional test: High vs. low analyst coverage. Cont'd

#### Table A15. Pre- and post- the financial crisis.

This table presents results for the pre-financial crisis subsample (1993-2007) in Panel A, and the post-financial crisis subsample (2010-2016) in Panel B. All variables are defined in Appendix A. All models include year and industry fixed effects. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively

	•	•	*		Panel A1.	Pre-crisis	(1)		· •	ž		
	Iss	sue	Freq	uency	Ra	nge	Ln(He	orizon)	Ор	tBias	Forece	ast_Miss
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Rep_Dum	0.109**		$0.112^{\dagger}$		$0.086^*$		0.086**		-0.123**		$-0.098^{*}$	
	(2.50)		(3.26)		(1.78)		(2.47)		(-2.49)		(-1.85)	
Rep_Index		$0.137^{**}$		$0.096^{**}$		$0.140^{**}$		$0.099^{**}$		-0.157**		-0.157**
		(2.48)		(2.27)		(2.31)		(2.30)		(-2.54)		(-2.39)
Observations	21,046	21,046	21,046	21,046	21,046	21,046	21,046	21,046	21,046	21,046	21,046	21,046
Pseudo /Adj. $R^2$	0.280	0.280	0.303	0.302	0.271	0.271	0.282	0.282	0.233	0.233	0.173	0.173
					Panel A2.	Pre-crisis	(2)					
	Асси	ıracy	Bad_	News	Good	News	Positive	Surprise	Negative	Surprise	Neutral	Surprise
Rep_Dum	$0.065^{\dagger}$		0.068		$0.093^{*}$		$0.092^{\dagger}$		<b>-</b> 0.144 [†]		0.043	
	(2.77)		(1.44)		(1.91)		(2.61)		(-3.76)		(0.80)	
Rep_Index		0.063**		0.122**		0.080		$0.105^{**}$		-0.167†		0.075
		(2.17)		(2.06)		(1.29)		(2.39)		(-3.51)		(1.13)
Observations	21,046	21,046	21,046	21,046	21,046	21,046	21,046	21,046	21,046	21,046	21,046	21,046
Pseudo /Adj. $R^2$	0.258	0.258	0.275	0.275	0.218	0.218	0.0514	0.0514	0.111	0.111	0.0805	0.0805
					Panel B1.	Post-crisis	(1)					
		sue	Freq	uency	Ra	nge		orizon)	Opt	Bias	Foreca	st_Miss
Rep_Dum	0.127**		$0.201^{\dagger}$		$0.166^{\dagger}$		$0.118^{**}$		-0.183†		$-0.207^{\dagger}$	
	(2.11)		(2.94)		(2.78)		(2.08)		(-2.92)		(-3.16)	
Rep_Index		0.057		0.089		$0.130^{*}$		0.052		-0.171**		-0.215†
		(0.81)		(1.12)		(1.83)		(0.78)		(-2.28)		(-2.75)
Observations	9,840	9,840	9,840	9,840	9,840	9,840	9,840	9,840	9,840	9,840	9,840	9,840
Pseudo /Adj. $R^2$	0.225	0.224	0.252	0.251	0.200	0.199	0.256	0.256	0.186	0.186	0.161	0.160

				]	Panel B2. 1	Post-crisis	(2)					
	Ассі	ıracy	Bad	News	Good	News	Positive	Surprise	Negative	Surprise	Neutral	Surprise
Rep Dum	0.103†		0.207†		-0.044		0.034		0.006		-0.124	
	(2.81)		(3.40)		(-0.64)		(0.63)		(0.10)		(-1.28)	
Rep Index		$0.085^*$		$0.162^{**}$		-0.128		-0.002		0.015		-0.030
		(1.93)		(2.23)		(-1.53)		(-0.04)		(0.22)		(-0.26)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
<b>Observations</b>	9,840	9,840	9,840	9,840	9,840	9,840	9,840	9,840	9,840	9,840	9,840	9,840
Pseudo $R^2/Adj$ . $R^2$	0.282	0.281	0.221	0.220	0.128	0.128	0.05	0.05	0.089	0.089	0.069	0.068

Table A15. Pre- and post- the financial crisis. Cont'd

#### Table A16. Robustness check. CEO Democratic ideology and MEF

This table presents tests of the association between CEO political ideology and management earnings forecast using the measure of a CEO's political ideology that captures Democratic affiliation. *Dem_Dum* is an indicator variable that equals one if a CEO donated more to the Democratic Party than to the Republican Party during her/his tenure. *Dem_Index* is the percentage of a CEO's support for the Democratic Party calculated as the number of cycles in which a CEO donates exclusively to the Democratic Party divided by her/his number of donation cycles in the sample period. Panel A reports results for the models of the association between CEO political ideology on one hand and *Issue, Frequency, Range, Ln(Horizon), OptBias, and Forecast_Miss,* on the other hand. Panel B reports results for the models of the association between CEO political ideology on one hand and *Accuracy, Bad_News, Good_News, Positive_Surprise, Negative_Surprise, and Neutral_Surprise* on the other hand. All other independent variables are defined in Appendix A. All models include year and industry fixed effects. T-statistics are computed using robust standard errors and reported in parentheses. [†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

	]	Panel A. C.	EO Politica	al ideology	and MEF	: measures	of CEO D	emocratic	ideology (	1)		
	Issue		Frequency		Range		Ln(Horizon)		OptBias		Forecast_Miss	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Dem_Dum	- 0.088 ^{**}		-0.008		-0.069		-0.063*		0.068		0.092*	
	(-2.18)		(-0.20)		(-1.61)		(-1.72)		(1.54)		(1.95)	
Dem_Index		-0.085		-0.006		-0.142**		-0.076		0.066		$0.121^{*}$
		(-1.40)		(-0.10)		(-2.23)		(-1.38)		(0.98)		(1.69)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
<b>Observations</b>	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951
Pseudo / Adj. $R^2$	0.257	0.257	0.279	0279	0.253	0.253	0.268	0.268	0.220	0.220	0.176	0.176
	]	Panel B. C	EO Politica	al ideology	and MEF	: measures	of CEO D	emocratic	ideology (2	2)		
	Accuracy		Bad News		Good_News		Positive_Surprise		Negative_Surprise		Neutral_Surprise	
Dem Dum	-0.025		-0.038		0.032		-0.077**		0.036		0.031	
—	(-0.70)		(-0.59)		(0.48)		(-2.20)		(0.65)		(0.57)	
Dem Index		-0.043*		-0.056		-0.016	. ,	-0.045		$0.074^{**}$		0.014
_		(-1.79)		(-1.33)		(-0.36)		(-0.86)		(2.00)		(0.18)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
<b>Observations</b>	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951	33,951
Pseudo / Adj. $R^2$	0.259	0.259	0.257	0.257	0.180	0.180	0.052	0.052	0.102	0.102	0.077	0.077

#### Table A17. Robustness check. Subsample of firms with management earnings forecasts.

This table presents the effect of the CEO political ideology on the management earnings forecasts by restricting the sample only to the guidance year. *Rep_Dum* is an indicator variable that equals 1 if a CEO donated more to the Republican party than to the Democratic party during her/his entire tenure. *Rep_Index* is the percentage of a CEO's support for the Republican Party calculated as the number of cycles in which a CEO donates exclusively to the Republican party divided by her/his number of donation cycles in the sample period. All other independent variables are defined in Appendix A. All models include year and industry fixed effects. T-statistics are computed using robust standard errors and reported in parentheses.[†], ^{**} and ^{*} denote significance at the 1%, 5%, and 10% levels, respectively.

	Frequency		Range		Ln(Horizon)		<b>OptBias</b>		Forecast_Miss		Accuracy	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Rep Dum	0.155†		$0.096^{*}$		0.004		-0.093*		-0.097**		$0.062^{\dagger}$	
1	(3.05)		(1.72)		(0.39)		(-1.82)		(-1.99)		(2.60)	
Rep Index		0.059		$0.123^{*}$		0.006	· /	-0.116*		$-0.174^{\dagger}$	· · · ·	$0.074^{**}$
· _		(0.94)		(1.76)		(0.45)		(-1.85)		(-2.92)		(2.50)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year & Ind. FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
<b>Observations</b>	11,988	11,988	11,988	11,988	11,988	11,988	11,988	11,988	11,988	11,988	11,988	11,988
Pseudo $R^2/Adj$ . $R^2$	0.286	0.285	0.146	0.146	0.069	0.069	0.153	0.153	0.113	0.114	0.278	0.278